

國立成功大學  
機械工程學系  
碩士論文

貝氏更新法應用於皮帶系統與不確定性壽命資料之設計整合

A Bayesian Based Updating Scheme for Belt-Pulley Systems Design  
with Censored Life Data

研究生：韓侑君  
指導教授：詹魁元博士

中華民國一百零二年六月

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Advisor: Dr. K.-Y. Chan

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## 摘要

皮帶系統為常見的動力傳輸系統。在此論文中，我們探討量產車輛中的皮帶傳動系統的設計，設計的主要目標為決定皮帶輪與張力輪的位置，以確保皮帶在靜態與動態行為下的表現均合乎要求，並考慮老化及不確定因素，使得皮帶系統在整個生命週期內均有良好的表現。然而，在皮帶設計過程中，仍然有許多困難尚未被解決。皮帶的結構為非均勻的複合材料，其中包含橡膠、鋼芯線體等。老化與不確定因素會使得皮帶特性難以預測，現有之皮帶設計分析均未將此因素列入考慮。若要準確的得知老化與不確定因素的影響，需要全面且大量的量測資料，但是在現實生活中，因為成本及資源有限，無法取得全面且大量的資料，所以重點式的量測為必要之方法。但是利用少量的量測點推估老化及不確定因素的影響，可能產生極大的誤差。所以需要一個可利用連續量測資料且可提供下一個重點量測參數的設計模式幫助我們進行皮帶系統設計。本論文中採用貝氏推估法進行皮帶系統之設計。貝氏推估法中的二項式推估，被用於評估皮帶系統在定時間下的可靠度表現；卜松推估，則被用於一段時間內的皮帶系統的破壞速率評估。在此論文中提出的設計模式中，先評估現有的資料的信心程度，而後再經由取樣重點參數提升信心程度。重點參數的選擇，取決於對有效限制式的敏感度分析結果。進一步利用蒙地卡羅過濾器去除偏頗之取樣。在本論文中提出的設計模式中，在不推估真實的老化與不確定因素模型的前提下，設計點必須滿足信心程度與可靠度要求，得到最化設計結果。在以一個數學範例演示此設計過程，並應用於一個工程皮帶系統設計問題上，得到一個考量老化與不確定因素的最佳設計。

**關鍵字:** 可靠度、不確定因素模型、貝氏定理、最佳設計、壽命資料、皮帶輪系統

# A Bayesian Based Updating Scheme for Belt-Pulley Systems Design with Censored Life Data

Student: I-Chun Han

Advisor: Dr. K.-Y. Chan

Department of Mechanical Engineering  
National Cheng Kung University

## ABSTRACT

Belt-pulley mechanism is commonly used in machinery and power transmission devices. In this thesis we investigate the use of tensioners of the belt-pulley mechanism inside a commercial vehicle. Design objective is to allocate the locations of pulleys and tensioners such that the static and dynamic behaviors of the entire system perform as desired throughout the entire life-time of the product. Design of the belt-pulley system suffers from the several issues that have not been fully addressed in the current literature. Most power transmission belts are semi-elastic transverse isotropic layered materials with steel core enhancements. Variation and deterioration of materials lead to uncertainty in materials that have not been accounted for in the current belt-related design problems. To obtain the precise material properties, extensive testing on various material properties are necessary. However, in reality the required measurement size is too large to provide abundant data; selective measurements are necessary due to time and other resource constraints. Unfortunately uncertainty models and the aging process can not be inferred accurately under few measurements. A design method that integrate sequential measurement data and also provide suggestions on additional data, whenever necessary, is needed. In this thesis we extend the Bayesian inference concept in the design of a more reliable belt-pulley system. Beta-binomial inference is used to estimate the reliability of a performance function given existing samples at a fixed time instant, an important tool to ensure product reliability at the initial state. Poisson-gamma inference is used to estimate the failure rate of a performance function given existing samples over a period of time. With the proposed method, we can first calculate the confidence with the current samples at hand and sequentially improve the confidence by adding samples. Addition samples are taken at the critical parameter decided by constraint activity and sensitivity analysis. An MCMC sample filter is applied to eliminate biased samples. The proposed design method will satisfy the confidence and reliability

targets without inferring the true uncertainty and the aging model with the fewest samples. A mathematical example is used to demonstrate this design method and the solution to the belt-pulley system design problem is then provided.

**Keywords:** reliability, uncertainty model, Bayesian theory, optimization, life data, belt-pulley systems.



# 誌謝

今天能有這本論文的產生，首先要感謝指導教授 詹魁元 老師在研究與待人處事上均給我很大的啟發，在研究過程中總是能將發現我太鑽牛角尖的思緒，並將我帶回主軸。在做人處事上，學習到老師的好脾氣，常常練習在事情多頭並行的時候能保持一個愉悅的心情，對人以鼓勵代替責難。在背後默默支持我的家人也是我心靈的支柱，父母給我一個無憂無慮的求學環境，家族的兄弟姐妹之間也常常互相打氣，心愛的姪子、姪女總是會傳可愛的聲音檔為我加油，讓我在夜深人靜的時候能有繼續下去的動力。因為有這些鼓勵與期待，讓我更努力完成學業，在外地一個人也不覺得辛苦。

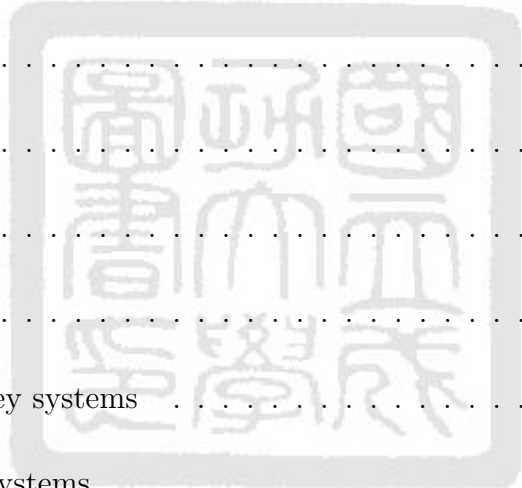
研究室的同學典運與東泰之間的情誼非常難得可貴，我們一起修過了艱難的最佳設計，度過無數個趕作業的晚上，謝謝你們的提醒與支援，讓我就算粗心大意也不至於壞事。在學術理論上，除了老師的指導，子頡學長是我隱藏版的導師，謝謝學長每次都犧牲睡眠時間幫我解答疑惑，就算人不在實驗室也會用SKYPE跟我討論問題。實驗室的學弟妹帶給我很多樂趣，我們一起慶生、一起烤肉，每天一起吃飯聊天。雖然我常常對你們生氣，其實還是非常感謝有你們的陪伴。肇余在半夜會傳來另類的鼓勵訊息，讓我緊繃的情緒能瞬間紓緩。我也不會忘記跟值榕半夜在實驗室打桌球、聊心事的回憶。也喜歡跟明證討論一些有關於未來的問題，雖然我們常常不同調，但是有人能討論未來與對人生的看法還是非常有趣。還有貼心的庭玉學妹，學妹總是把實驗室的雜事規畫得井井有條，另外，親手做的卡片真的讓我非常感動。實驗室的人員眾多，煜駿、佑安、彥彬、承勳、米約瑟與祐晨，無法一一詳述，感謝大家一起組成這個美好的系統最佳化實驗室。也謝謝設計組的大家，一起出遊團拍畢業照，幫我把碩士服從海裡撿回來，很開心能待在這麼熱情的設計組。

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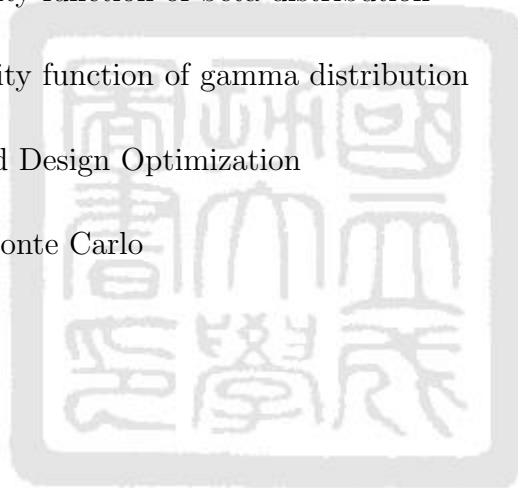
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# List of Symbols

$a$	parameter of gamma distribution
$a_k$	acceptance probability of MCMC filter
$A$	cross section
$b$	parameter of gamma distribution
$CR_R (CR_f)$	confidence range for time-dependent (time-independent) constraint
$CB_R (CB_f)$	confidence bound for time-dependent (time-independent) constraint
$CR_t (CL_t)$	confidence range target for time-dependent (time-independent) constraint
$d$	deterministic design variables
$D_u$	uncertain design variables with known distributions
$E$	Young's modulus
$EI$	bending stiffness
$EA$	longitudinal stiffness
$f$	friction force
$g$	deterministic constraint
$g_R$	reliability constraints with reliability target
$g_B$	Bayesian reliability constraints with a reliability target and a confidence range target
$g(t)$	time-dependent deterministic constraint
$g_R(t)$	time-dependent reliability constraints with reliability target
$g_B(t)$	time-dependent Bayesian reliability constraints with a reliability target and a confidence range target

$I$	moment of inertia
$k$	parameter of Poisson distribution; expected number of failure
$M_i$	torque on $i$ th pulley
$n$	normal force
$N$	number of samples
$p$	probability of successful outcomes
$\mathbf{p}$	deterministic parameters
$\mathbf{P}_u$	uncertain parameters with known distributions
$\mathbf{P}_s$	uncertain parameters with samples
$P_f$	probability of failure
$P_{i,j}$	transition probability from $i$ to $j$
$q$	proposal distribution
$Q$	shear force
$r$	number of successful outcomes
$R_a$	reliability of aging process
$R_t (P_t)$	reliability target for time-dependent (time-independent) constraint
$r_i$	radius of $i$ th pulley
$T$	tension force
$\alpha$	parameter of Beta distribution
$\beta$	parameter of Beta distribution
$\theta$	angle between a belt and x-axis
$\kappa$	curvature of a belt

$\lambda$	parameter of Poisson distribution; average failure rate
$\mu_{\mathbf{D}_u}$	mean of design variables
$\mu_{\mathbf{P}_u}$	mean of design parameters
$\pi$	target distribution
$\phi_{\text{beta}}$	probability density function of beta distribution
$\phi_{\text{gamma}}$	probability density function of gamma distribution
$\phi_i$	wrapping angle of $i$ th pulley
$\Phi_{\text{beta}}$	cumulative density function of beta distribution
$\Phi_{\text{gamma}}$	cumulative density function of gamma distribution
RBDO	Reliability-Based Design Optimization
MCMC	Markov chain Monte Carlo



# Chapter 1 Introduction

In chapter 1, a general belt-pulley system model will first be introduced including the basic background introduction in section 1.1 and performance analysis of belt-pulley system in section 1.2. However, some obstacles occur when applying these belt-pulley system models in practice. Therefore, the motivation and the research objectives of this thesis are addressed in section 1.3. In section 1.4, we will introduce the structure of this thesis.

## 1.1 Background of belt-pulley systems

A belt-pulley system is an important transmissive mechanism in modern machinery. Figure 1.1 shows a general belt-pulley system with a tensioner. Pulley 1 is a driving pulley, pulley 2 is a driven pulley and pulley 3 serves as a tensioner.  $\phi_i$  is wrapping angle of the  $i$ th pulley. In this system, Torque transmits from pulley 1 to pulley 2 and the position of tensioner influences a transmissive efficiency. With a fixed belt length, The lower the position of tensioner is, the larger the tension exerts on belt and the bigger wrapping angle will be.

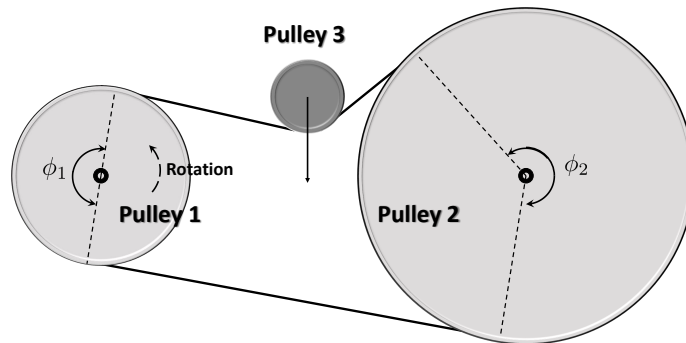


Figure 1.1: Two pulleys belt system

Belt-pulley systems are widely applied in our daily life. Some of them are used to deliver power; while others serve as transmission belt. Belt-pulley systems exist in sewing machine, fitness equipment and various vehicle transmission system as shown in Figure 1.2. The focus of this thesis is on the automobile transmission model. This belt-pulley system usually equips



a tensioner to maintain sufficient tension to ensure transmitting torque from engine to gears is acceptable. The position of the tensioner would effect transmissive efficiency as well as power efficiency. When the tensioner provides overloading tension, the transmit efficiency might be high but the power efficiency would be low, and vis versa. Therefore, a proper position balancing transmit efficiency and power efficiency is an important issue.

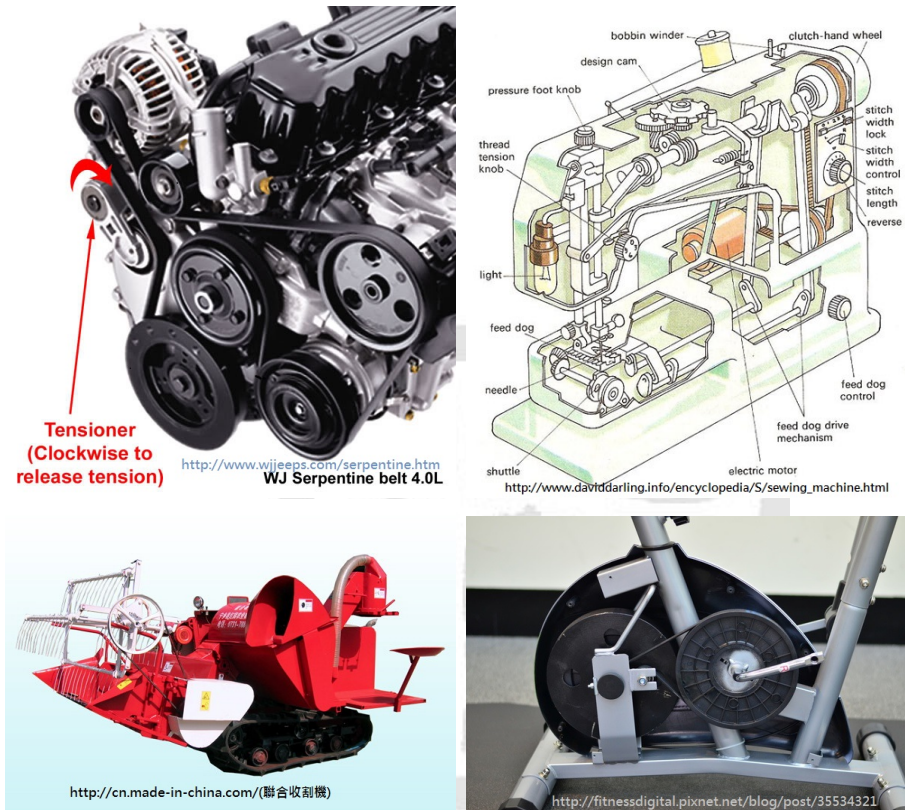


Figure 1.2: Applications of belt-pulley system

## 1.2 Analysis of belt-pulley systems

After providing basic belt-pulley systems concept, let us look at the standard design guide-lines of belt-pulley systems [1,2]. These guide-lines usually starts with a general two pulleys system as shown in Figure 1.3. A flat belt is treated as a string in traditional analysis of a belt-pulley system. Pulley1 is a driving pulley and pulley 2 is a driven one. A part of flat belt between two pulleys, span, is a straight line and tangents to pulleys. Contact points would be obtained simply by analyzing a geometric outline of belt-pulley system.  $T_i$  is a torque on  $i$ th pulley ,  $r_i$  is the radius of the  $i$ th pulley and  $P_1$  and  $P_2$  are tension forces. The relationship follows [1]:

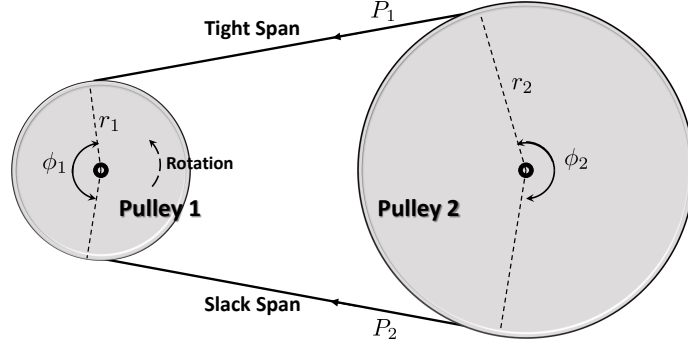


Figure 1.3: Two pulleys belt system

$$T_i = (P_1 - P_2) \times r_i \quad (1.1)$$

When a belt transmits torque from one pulley to another, two parts of span have different tensions depending on the direction of motion. One of span is exerted larger tension force; it is called the tight span. On the contrary, another span with smaller tension is called the slack span. In order to analyze the force applied on different part of a belt, a small element of belt is studied. For small angel  $d\theta$ , tension force  $P$  is the friction coefficient  $f$  times the normal force  $dN$ . From Figure 1.4, we know  $dN = P \times \sin d\theta/2$  and for a small  $\theta$ ,  $\sin d\theta/2 \approx d\theta/2$ . Therefore

$$dP = f \times dN, \quad dN = 2(P \times d\theta/2) = pd\theta \quad (1.2)$$

From Equation(1.2), we can then yield:

$$\frac{dP}{P} = fd\theta \quad (1.3)$$

Integrating both sides of Equation(1.3) with upper and lower bounds of  $dP$  be  $P_2$  and  $P_1$  separately and the range of  $\theta$  is 0 to  $\phi$ .

$$\int_{P_2}^{P_1} \frac{dP}{P} = \int_0^\phi fd\theta \quad (1.4)$$

The final governing equation yields:

$$\frac{P_1}{P_2} = e^{f\phi} \quad (1.5)$$

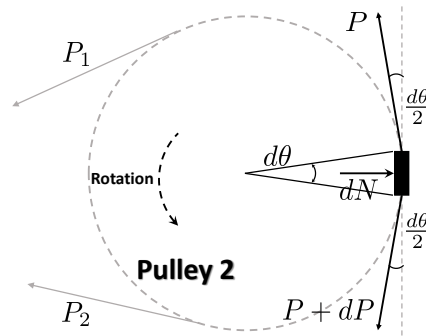


Figure 1.4: A free body diagram of belt element

However, these analyses are not sufficient for modern belts with complex structures. Figure 1.5 is a standard cross section of a belt constructed by multiple layers. The performance analysis in belt-pulley systems is therefore a complex problem due to the compound layers of its structure and unpredictable aging performance of rubbers. Steel strings and fabrics within belt provide additional resistance to be extended and curved. The resistance of extending is called longitudinal stiffness and the resistance of curving is called bending stiffness. In existing belt analysis, These two resistances are not taken into consideration and neither do the material properties.

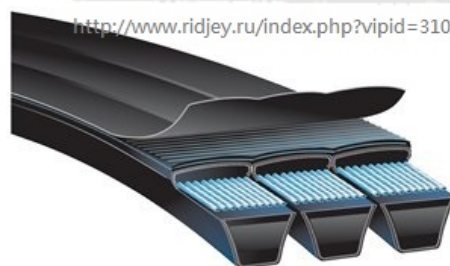


Figure 1.5: A cross section of belt

### 1.3 The need of design method with uncertainty data

Material and physical properties such as bending stiffness and longitudinal stiffness would influence wrap angles as in Figure 1.6 and consequently affect the transmissive efficiency. Therefore,

providing a proper value of these properties is important. These properties can no longer be

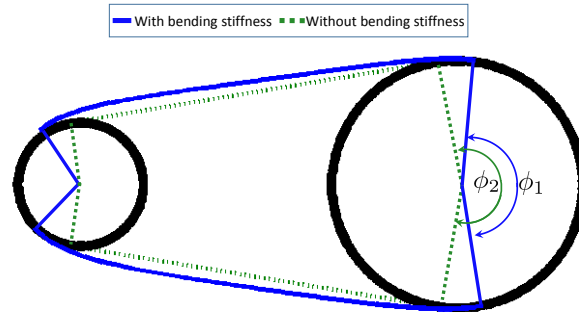


Figure 1.6: The effects of bending stiffness

assumed constant. Furthermore, uncertainties in manufacture process make material properties vary from one to another. For example:

- The Young's modulus  $E$  is non-homogeneous due to the compound layers of belt.
- The moment of inertia  $I$  is another complicated problem to be addressed. Both the curvature of span and a center of rotation are changing and the shape of belt cross section is special. Hence, moment of inertia  $I$  is difficult to calculate in practice.
- Cross section  $A$  of a belt would vary from belt to belt due to tolerance in manufacturing process.

All variations of material properties and geometric dimensions are **uncertainties**. Uncertainty would lead to unpredictable effects on system and makes system perform against original design. Statistically, if we know the types of uncertainty, we can model it to modify our design. There are many design method exist to address this problem by assuming the types of uncertainty. We will introduce them in section 2.3. Unfortunately, the true type of uncertainty is never know for us. All we can do is inferring the type uncertainty based on existing samples. Because of the lack of time and limited resource, we generally do not have enough sample points to verify a type of uncertainty with enough confidence. A limit number of sample points might mislead us and comes to an wrong conclusion. Some methods have been developed to tackle

this task. Sample points are not use to infer the true types of uncertainty directly but infer the confidence level of these sample points.

Furthermore,uncertainty may change over time. Even though a system work well as new, it might deteriorate very fast because of degeneration of components. For example, a belt might be stretched more easily after two years; therefore, the requires target tension cannot be reached so the belt-pulley system will eventually fail. These aging phenomenons would affect a performance of a system.

Taking the uncertainty and aging effects into design consideration is needed. If we consider variation and deterioration in early design stage, we do not only ensure a performance at initial time but also control the aging rate. We can infer performance and aging rate by sample life data which contain time-independent properties and time-dependent one.

In this thesis, we will:

- Treat uncertainty as unknown and simply record measurements of parameters.
- Use Bayesian inference to analyze how uncertainty influent systems with life data.
- Develop a design scheme that adding Bayesian analyses with life data.

## 1.4 Thesis organization

In chapter 2, we review some design methods according to data quality. When the type of data is ideal, Some deterministic optimal design method and model are introduced in section 2.2. For these data which are not perfect but provide a large amount of information to us, the reliability based design methods is required and they will be introduced in section 2.2. As the quality of information decreasing, reliability based design method need to be modified and new design method will be introduced in section 2.3. Furthermore, aging effects are taken into consideration so some methods focus on analyzing failure rate developed. Basic concepts of

them are introduced in section 2.4. In chapter 3, we want to use Bayesian inference to analyze the performance and aging process of a system. First, In section 3.1, A brief introduction of Bayesian inference are shown including Bayes theorem in subsection 3.1.1 and two forms of Bayesian inference; empirical and hierarchical Bayesian inference in subsection 3.1.2 and 3.1.3. Then, In section 3.2, illustrating what kind of prior in Bayesian inference is proper according to what kind of data we have. Demonstration of Bayesian method applying reliability estimation is shown in section 3.3. In chapter 4, the analysis which is demonstrated in section 3.3 is adding to design process. The flowchart of proposed design method present in section 4.1 and each step in the flowchart will be introduced in the following section 4.2 to 4.4. In chapter 5, we apply proposed design method to a mathematical example and a belt-pulley system design case. A conclusion and future work are shown in chapter 6.



# Chapter 2      Review of Belt-Pulley System

## Analysis and Design Methods

### 2.1 Belt-pulley system performance analysis

There are two basic theories in the belt-pulley system analysis, namely shear theory [3–5] and creep theory [5–7]. Creep theory is the original theory to analyze the performance of a belt-pulley system and it is usually applied to soft belt such as those made of leather only. Soft belt could perfectly fit the profile of pulleys and the span between any two pulleys can reasonably be approximated as a straight line. However, belts we use nowadays are not leather belts anymore; instead, they are made of layers of rubber, steel cores, and artificial textiles with the shape of V. Features of V-belts differ much from leather belts. V-belt cannot wrap around the pulley profiles perfectly and the span between two pulleys is a curve, not a straight line. Shear theory has therefore been developed to deal with this new kind of belts. However, the calculation in shear theory is too complicated to widely be applied.

In fact, these phenomena happen in V-belts mainly because of a material property of a belt called bending stiffness. Some researchers try to take bending stiffness into consideration when analyzing belt-pulley system with creep theory. Analysis of belt-pulley system in this thesis is mainly based on creep theory considering bending stiffness by Parker et al [8,9].

We using a two-pulley belt system to illustrate the analysis process that can then be extended to general cases with multiple pulleys [8]. A belt is separated into three parts, namely span, adhesion and sliding. Span represents part of a belt that dose not in contact with any pulley. Adhesion zone is the part of a belt that moves in the same velocity with the pulley. Finally, in the sliding zone, a velocity difference exists between sliding pat of belt and pulley. In Figure 2.1, pulley 1 is the driver with driving torque  $M_1$ .  $M_2$  is a loading exerting on pulley 2.  $\alpha_i$  is sliding angel,  $\beta_i$  is adhesion angel on  $i$ th pulley. Behaviors of a belt are governed by equations with different physical characteristics for different part of a belt. Material constants

in these governing equations are different at each zone.

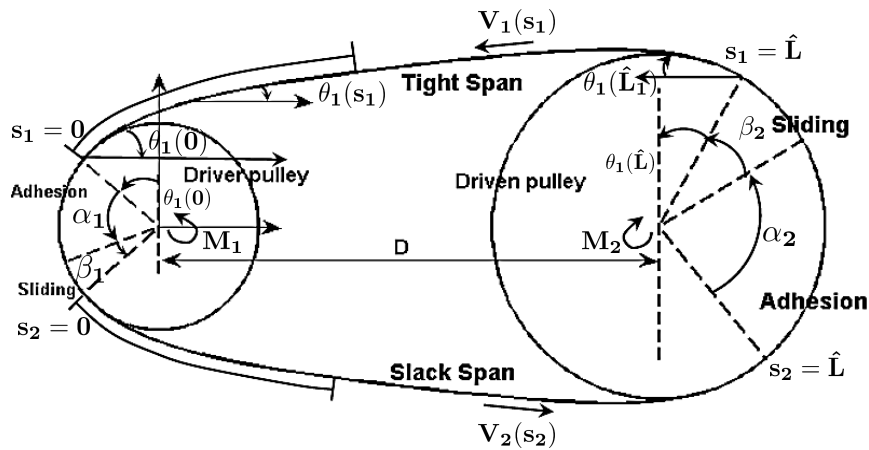


Figure 2.1: Two pulleys belt system

Derivation of the governing equations begins from the free body diagram of a stretched belt section in Figure 2.2. A small section of a stretched belt is exerted by a tension force  $T(s)$ , shear force  $Q(s)$  and a bending moment  $M(s)$ .  $n(s)$  and  $f(s)$  are the normal and the friction forces between a belt and a pulley.

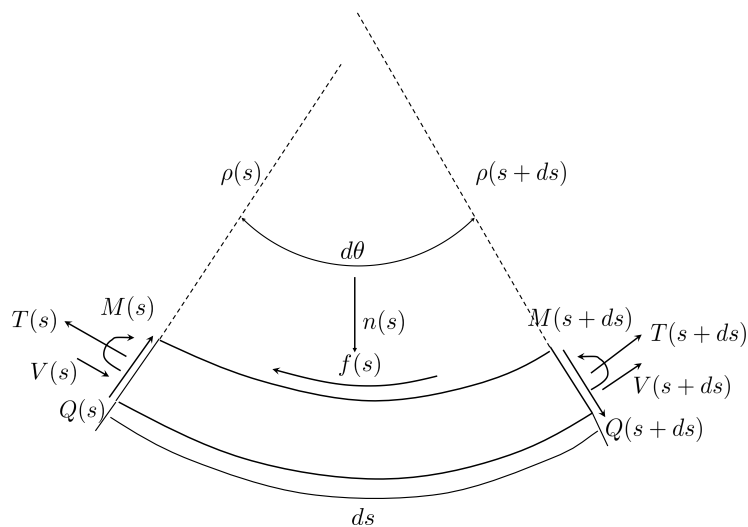


Figure 2.2: Free body diagram



This small section of a belt is treated as a moving Euler-Bernoulli beam and the Euler-Bernoulli theory requires :

$$M = EI\kappa \quad (2.1)$$

where  $EI$  is the bending stiffness,  $E$  is the Young's modulus,  $I$  is the moment of inertia, and  $\kappa$  is the curvature of the beam which also be defined as changing rate of inclination angle  $\theta$ ,  $\kappa = d\theta/ds$ . Since the beam is in a steady motion, forces acting on it should balance:

- The balance of angular momentum with respect to the center of mass of the small section results in:

$$dM - Qds = 0, \quad Q = dM/ds = EI(d\kappa/ds) \quad (2.2)$$

- The balance of linear momentum in tangential direction results in:

$$dT - fds = GdV - Qd\theta, \quad G = m(s)V(s) = \text{constant} \quad (2.3)$$

where  $m(s)$  is the mass density per unit length, For steady state, mass flow is conserved so  $G = \text{constant}$ .

- The balance of linear momentum in normal direction results in:

$$(T - GV)d\theta - dQ = nds \quad (2.4)$$

From Equation(2.1) to Equation(2.4), we can yield the following two governing equations:

$$(T - GV)' + EI\kappa\kappa' = f \quad (2.5a)$$

$$(T - GV)\kappa - EI\kappa'' = n \quad (2.5b)$$

Equations(2.5) guide the behavior of the whole belt but a little different in each part. For example, the friction force in Equation(2.5a) is different for sliding zone and adhesion zone because of the difference in static and kinetic frictional coefficient. Furthermore, for the span part, the belt does not contact with any pulley so the contact forces  $f$  and  $n$  are zero. Equation(2.5) is then transferred into:

$$(T - GV)' + EI\kappa\kappa' = 0$$

$$(T - GV)\kappa - EI\kappa'' = 0 \quad (2.6)$$

Equation(2.5) is the main governing equations in belt-pulley system analysis. Other geometric boundaries are required when solving Equation(2.5). These boundaries are there to ensure the belt exactly contacts with pulleys and they are orthogonal at the points of contacts. The definition of boundary is shown in Figure 2.1.

Let  $T - GV = W$ . Seven differential equations and seven boundaries are shown as following:

- Physical differential equations:

- $dW/ds = EI\kappa\kappa' - f$ , derived from Equation(2.5a).

- $d\kappa/ds = (f - W)/\kappa \times EI$ , derived from Equation(2.5a).

- $d^2\kappa/ds^2 = (W - n)/EI$ , derived from Equation(2.5b).

- Geometric differential equations:

- $d\theta/ds = \kappa$ , it is the definition of curvature.

- $dx/ds = \cos \theta$ ,  $x$  is a position on x-axis.

- $dy/ds = \sin \theta$ ,  $y$  is a position on y-axis.

- $d\hat{L}/ds = 0$ ,  $\hat{L}$  is the length of span in steady motion and it is a constant.

- Boundary conditions:

- $W(0) = T(0) - GV(0)$ , the initial value of span closed to pulley<sub>*i*</sub>

- $\kappa(0) = 1/R_i$ , at points of contacts, the curvature of belt is equal to which of pulleys.

- $\kappa(\hat{L}) = 1/R_{i+1}$

- $x(0) = R_i \sin(\theta(0))$ , ensuring the belt is orthogonal to pulleys.

- $x(\hat{L}) - D = R_{i+1} \sin(\theta(\hat{L}))$

- $x(0)^2 + y(0)^2 = R_i^2$ , ensuring the belt contact with pulleys.

- $x(\hat{L})^2 + y(\hat{L})^2 = R_{i+1}^2$

By solving this boundary value problem, we can yield  $W$ ,  $\theta$ ,  $\kappa$ ,  $\kappa'$ ,  $x$  and  $y$ . Detail of solving process will be introduced in section 5.2.

## 2.2 Design methods with ideal data

After describing the analysis of belt-pulley system, let us switch topic to general design methods of engineering product. Theoretically, we can improve the quality and performance of an engineering product by optimization. We set an objective function ( $f$ ) and several constraints ( $g_i$ ) and then minimize the objective function subject to these constraints within the design upper ( $\mathbf{d}^U$ ) and lower bounds ( $\mathbf{d}^L$ ). When parameters ( $\mathbf{p}$ ) are constants, we can yield an optimal design ( $\mathbf{d}^*$ ) by solving the optimization problems posed as :

$$\begin{aligned} & \min_{\mathbf{d}} f(\mathbf{d}, \mathbf{p}) \\ \text{s.t. } & g_i(\mathbf{d}, \mathbf{p}) \leq 0, i = 1 \sim n \\ & \mathbf{d}^L \leq \mathbf{d} \leq \mathbf{d}^U \end{aligned} \quad (2.7)$$

## 2.3 Design methods with abundant data

In reality, uncertainty exists in manufacturing process, in material properties, and in the variation of environment, some parameters  $\mathbf{P}$  may not be constants anymore. Parameters would vary in a small range. These unpredictable effects are called uncertainty. When design process involves uncertainty, the original optimal process is transferred into a probabilistic optimal design format [10–12] as in Equation(2.8) which are referred to as reliability-based design optimization (RBDO). Equation(2.8) is a generalized single-objective probabilistic formulation with random design variables  $\mathbf{D}$ , random parameters  $\mathbf{P}$ , deterministic design variables  $\mathbf{d}$  and deterministic parameters  $\mathbf{p}$ . The objective  $f$  is a function of the deterministic quantities and the mean values of random quantities in the formulation. The feasible space of  $\mathbf{d}$  subject to all constraints in  $\mathcal{K}$  is  $\mathcal{F}$ . The reliability of a design is defined as the probability of satisfying constraints, denoted as  $1 - P_f$

$$\begin{aligned} & \min_{\mu_{\mathbf{D}}, \mathbf{d}} f(\mu_{\mathbf{D}}, \mu_{\mathbf{P}}, \mathbf{d}, \mathbf{p}) \\ \text{Pr}[g_j(\mathbf{D}, \mathbf{P}, \mathbf{d}, \mathbf{p}) > 0] & \leq P_{f,j} \quad \forall j \in \mathcal{K} \end{aligned} \quad (2.8)$$

Solving the probabilistic constraint require an additional analysis to the conventional opti-

mal design process. Therefore, the solution complexity and the computational cost are increasing. Standard approaches in solving this probabilistic constraints include: the first/second order reliability method (FORM/SORM) [13–15], adaptive importance sampling [16], advance mean value [17], and its hybrid variant [18], sequential optimization and reliability assessment [19], and single-loop method [20]. Furthermore, many methods have been proposed to enhance numerical efficiency and stability in solving RBDO problems [21–23].

## 2.4 Design methods with inadequate data

Applying the probabilistic methods mentioned in section 2.3 requires the distribution of the uncertainty or the aging process be known a priori. Unfortunately, it is impractical to know the exact distributions of uncertainty or the aging process; the conventional methods, are therefore limited in its industrial practice.

Instead of have the distribution information, we usually have only limited measurement samples of these uncertainty. These sample points are measurements drawn from specific distribution and the type of distribution is unknown to us. Some researchers use sample points to infer the unknown distribution of uncertainty. The inference generates large errors when the sample points are inadequate or type of distribution we assumed is improper [24–26]. Other researchers analyze reliability without inferring the distributions of uncertainty; they use the concept of confidence in design, such as possibility-based design optimization (PBDO) [27,28] based on possibility theory [29–34], evidence-based design optimization (EBDO) [35] based on evidence theory [36–38], and Bayesian RBDO [39–41] based on Bayes theory [42–45].

In Bayesian RBDO, the analysis of confidence levels based on data is added into the reliability-based optimization process. The generalized Bayesian optimization model can be

expressed as:

$$\begin{aligned}
& \min_{\mu_{\mathbf{D}_u}, \mathbf{d}} f(\mu_{\mathbf{D}_u}, \mathbf{d}, \mathbf{P}_s, \mu_{\mathbf{P}_u}, \mathbf{p}) \\
& \text{s.t } g_i = g^i(\mathbf{d}, \mathbf{p}) \leq 0 \\
& g_R = \Pr[g^i(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}) \leq 0] \geq R_t \\
& g_B = \Pr [\Pr[g^i(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{P}_s, \mathbf{p}) \leq 0] \geq R_t] \geq CR_t
\end{aligned} \tag{2.9}$$

where

$\mathbf{d}$  : deterministic design variables

$\mathbf{D}_u$  : uncertain design variables with known distributions

$\mathbf{p}$  : deterministic parameters

$\mathbf{P}_u$  : uncertain parameters with known distributions

$\mathbf{P}_s$  : uncertain parameters with samples

$g$  : deterministic constraint

$g_R$  : reliability constraints with reliability target

$g_B$  : Bayesian reliability constraints with a reliability target and a confidence range target

The main idea of Bayesian inference is that a posterior distribution is proportional to the product of likelihood and prior distribution [39]. The prior is an initial guess of distribution about on uncertainty, the posterior is an inferred distribution according to some observations, and the likelihood can be treated as a weighting between posterior and prior. The type of prior would effects final result of posterior so some researchers try to figure out how to choose the proper prior when applying Bayesian inference [46].

Computational operation in Bayesian inference sometimes could be very expensive. Conjugate prior could help us get arid of expensive computation. We could obtain a posterior distribution by changing the parameters of prior distribution instead of by multiple integration. Different types of conjugate prior we choose could address different kind of problems. For instance, the Binomial-Beta inference choosing beta distribution as a prior distribution and

Binomial distribution as likelihood is mainly use to address problem with pass/fail data at a specific time instant [39]. Moreover, the Poisson-Gamma inference setting a Gamma prior and Poisson likelihood is frequently applied to calculate the probability of failure given a failure rate and an expected failure number. Bayesian inference could deal with both time-independent and time-dependent problem by choosing different conjugate priors.

## 2.5 Design methods with life data

A good product is not only reliable but also durable; that means low probability of failure at the initial time and low failure rate throughout the life span. Figure 2.3 shows a typical failure probability curve with respect to time. The solid curve is the current design and the dash line is the design with improved reliability, which has low probability of failure and small the increment of failure rate.

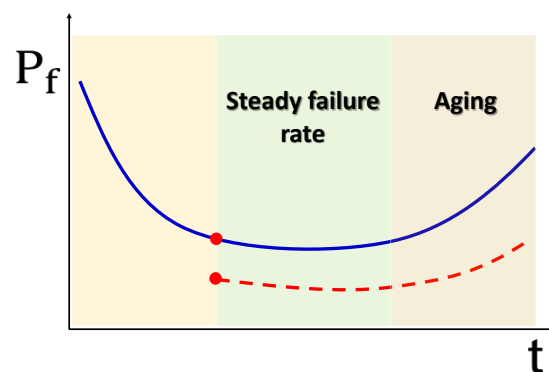


Figure 2.3: Bathtub curve in product reliability

After including time-factor in design, the RBDO and Bayesian RBDO are transferred into time-dependent RBDO and time-dependent Bayesian RBDO. The general mathematical formulation can be expressed as:

$$\begin{aligned}
& \min_{\boldsymbol{\mu}_{\mathbf{D}_u}, \mathbf{d}, \mathbf{P}_s, \boldsymbol{\mu}_{\mathbf{P}_u}, \mathbf{p}} f(\boldsymbol{\mu}_{\mathbf{D}_u}, \mathbf{d}, \mathbf{P}_s, \boldsymbol{\mu}_{\mathbf{P}_u}, \mathbf{p}) \\
& \text{s.t. } g_i = g^i(\mathbf{d}, \mathbf{p}) \leq 0 \\
& g_R = \Pr[g^i(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}) \leq 0] \geq R_t \\
& g_B = \Pr[\Pr[g^i(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{P}_s, \mathbf{p}) \leq 0] \geq R_t] \geq CR_t \\
& g_i(t) = g_t^i(\mathbf{d}, \mathbf{p}, \mathbf{p}(t)) \leq 0 \\
& g_R(t) = R_a(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}, \mathbf{D}_u(t), \mathbf{P}_u(t), \mathbf{p}(t)) \geq P_t(t) \\
& g_B(t) = \Pr[R_a(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}, \mathbf{D}_u(t), \mathbf{P}_u(t), \mathbf{p}(t)) \geq P_t(t)] \geq CL_t(t)
\end{aligned} \tag{2.10}$$

where

$\mathbf{d}$  : deterministic design variables.

$\mathbf{D}_u$  : uncertain design variables with known distributions.

$\mathbf{D}_u(t)$  : time-dependent uncertain design variables with known distributions.

$\mathbf{p}$  : deterministic parameters.

$\mathbf{p}(t)$  : time-dependent deterministic parameters.

$\mathbf{P}_u$  : uncertain parameters with known distributions.

$\mathbf{P}_u(t)$  : time-dependent uncertain parameters with known distributions.

$\mathbf{P}_s$  : uncertain parameters with samples.

$\mathbf{P}_s(t)$  : time-dependent uncertain parameters with samples.

$g$  : deterministic constraint.

$g_R$  : reliability constraints with a reliability target  $R_t$ .

$g_B$  : Bayesian reliability constraints with a reliability target  $R_t$  and a confidence range target  $CR_t$ .

$g(t)$  : time-dependent deterministic constraint.

$g_R(t)$  : time-dependent reliability constraints with a reliability target  $P_t$ .

$g_B(t)$  : time-dependent Bayesian reliability constraints with a reliability target  $P_t$  and a confidence range target  $CL_t$ .

$R_a$  : the reliability of aging process which is inferred based  $\mathbf{p}$ ,  $\mathbf{p}(t)$ ,  $\mathbf{P}_u$ ,  $\mathbf{P}_u(t)$ ,  $\mathbf{P}_s$  and  $\mathbf{P}_s(t)$ .

$R_t$  and  $P_t$  : the reliability target for time-dependent and time-independent cases.

$CR_t$  and  $CL_t$  : the confidence range target for time-dependent and time-independent cases.

Three basic approaches have been proposed to deal with time-dependent reliability-based optimization, namely the extreme performance approach, the first-passage approach, and the composite limit state approach. In extreme performance approach, response surfaces of constraints and the objective function are built to reach a maximal improvement of a performance criteria, i.e., the accuracy. Although the performance of a response surface is updated by adding new samples as shown in [47], these approaches isolate time factor instead of considering time-varying characteristics. In other words, the aging model is known. Another widely applied method is the first-passage approach. The general concept of the first-passage approach uses the up-crossing rate to evaluate the probability of failure over a period of time when initial failure occurred [47,48]. However, the up-crossing rate is hard to calculate so there are some methods developed to reduce computational efforts [48]. The third method to address a time-dependent RBDO is the composite limit state approach [49] that divides a time interval into finite time periods and transfers a time-dependent problem into a time-independent equivalence [47]. Note that the extreme performance approach, the first-passage approach, and the composite limit state approach can be applied to both inferences with abundant data and inadequate data.

Other relevant approach about time-dependent reliability analysis focus inferring the failure rate. For example, Colombo et al. presented nonparametric estimation of time-dependent failure rates with life data [50]. Semi-parametric bathtub-curve failure rate was proposed by Ho [51]. However, nonparametric estimation and Semi-parametric of failure rate require a abundant of data, but sometimes we are not able to get so much data to construct model accurately.



# Chapter 3 Bayesian Inference in Reliability

## Analysis with Mixed Data Types

The main goal of this research is to systematically incorporate measured data within the design process. The central idea behind this process is the Bayesian updating scheme. In this chapter, Bayes theorem would be introduced in section 3.1 followed by two types of conjugate prior to address time-independent and time-dependent problem in section 3.2. The procedure of reliability estimation is presented step by step in section 3.3. Finally, an constraints reliability estimation example is demonstrated in section 3.4.

### 3.1 Introduction of Bayesian inference

Uncertainties are omnipresent in design process and designers usually assume they know the underlying distributions of uncertainties. However, the underlying distribution is never know to us. When an assuming distribution differs from the true underlying distribution, inference would lead to wrong results. Even if the type of distribution assumed is correct, inferring underlying distribution with limit number of samples might be misled by bias samples. Therefore, a confidence range of these data is calculated as an expression of the degree of trust on measured data. Bayesian inference is the approach based on Bayes theorem to estimate the confidence range of data. In what follows, we will talk about Bayes theorem first.

#### Bayes theorem

Bayes theorem is the rule based on conditional probability as in Equation(3.1):

$$\Pr(A|B) = \frac{\Pr(A \cap B)}{\Pr(B)} \quad (3.1)$$

where  $\Pr(A|B)$  is the probability of an event A happening given event B,  $\Pr(B)$  is the probability of an event B, and  $\Pr(A \cap B)$  is the joint probability of both events A and B. Furthermore,  $\Pr(A \cap B) = \Pr(B|A) \Pr(A)$  is known as **multiplication rule** can transfer Equation(3.1) into Equation(3.2):

$$\Pr(A|B) = \frac{\Pr(B|A) \Pr(A)}{\Pr(B)} \quad (3.2)$$

Define  $\Pr(A|B)$  as the posterior probability,  $\Pr(B|A)$  as the likelihood, and  $\Pr(A)$  as the prior probability. According to Equation(3.2),  $\Pr(A|B) \propto \Pr(B|A) \Pr(A)$ . This is the basic concept of Bayes inference, which states: the posterior probability is in proportion to the product of the prior and the likelihood.

If events A and B follow binomial process, the formula in Equation(3.2) could be expressed in another form. We denote  $\Pr(A^c)$  as the complement probability of the event A. The probability of the event B happening can be denoted as Equation(3.3)

$$\Pr(B) = \Pr(B \cap A) + \Pr(B \cap A^c) \quad (3.3)$$

Applying multiplication rule to each joint probability, we can rewrite Equation(3.2) as:

$$\Pr(A|B) = \frac{\Pr(B|A) \Pr(A)}{\Pr(B)} = \frac{\Pr(B|A) \Pr(A)}{\Pr(B|A) \times \Pr(A) + \Pr(B|A^c) \times \Pr(A^c)} \quad (3.4)$$

In more general cases with multiple events, we then have:

$$\Pr(B) = \sum_{j=1}^k \Pr(B|A_j) \Pr(A_j) \quad (3.5)$$

Conditional probability of multiple events results in Equation(3.6).

$$\Pr(A_j|B) = \frac{\Pr(B|A_j) \Pr(A_j)}{\sum_{j=1}^k \Pr(B|A_j) \Pr(A_j)} \quad (3.6)$$

Equation(3.6) is the general form of **Bayes theorem**. The form also express the basic idea of Bayes inference:

$$\textit{Posterior Probability} \propto \textit{Prior Probability} \times \textit{Likelihood}$$

Note that when applying the Bayes inference, the prior probability,  $\Pr(A_j)$ , can be treat as the bound of previous known information, the posterior,  $\Pr(A_j|B)$ , can be treated as the inferred result based on known information, and the likelihood,  $\Pr(A_j|B)$ , presents the similarity between the prior and the posterior.

## 3.2 Prior selection based on data types

The types of prior distribution and of likelihood result in different type of posterior distributions. In some special cases, the types of prior and posterior distributions are the same; these priors are

called conjugate priors. The biggest advantage of conjugate priors is the parameter calculation with known distributions. Beta-binomial and Poisson-gamma are the common used conjugate priors. We will introduce the use of beta-binomial to infer time-invariant uncertainty data in Section 3.2.1 and the use of Poisson-Gamma to infer time-variant life data in Section 3.2.2 3.2.2, respectively.

### 3.2.1 Time invariant measurement data

When the prior distribution follow a beta distribution and the likelihood follow the binomial process, the posterior distribution will also be a beta distribution, a standard process called beta-binomial inference. Beta-binomial inference usually applies in binary tests such as pass/fail test at a time instant.

#### Beta Distribution

Beta distribution is a continuous distribution defined on the interval  $[0,1]$ , with two positive parameters  $\alpha$  and  $\beta$

as in Equation (3.7) where  $\Gamma(\cdot)$  is the Gamma function.

$$f(p|r) = \frac{\Gamma(\alpha + \beta + N)}{\Gamma(\alpha + r)\Gamma(N - r + \beta)} p^{(r+\alpha-1)}(1 - p)^{(N-r+\beta-1)} \quad (3.7)$$

Taking the average death rate of a disease for instance,  $p$  represents the average death rate and  $f(p|r)$  is the probability of the expected death rate occurring given the observed death number. Note that beta distribution with parameters (1,1) is an uniform distribution, that means all the probability of expected probability is the same.

#### Binomial Distribution

Binomial distribution is frequently used as to model the probability of a specific number of marked sample occurring given a probability of expected average in a binary experiment. Let  $r$  be the number of success,  $N$  be the total sample number, and  $p$  be the average probability of success in Equation (3.8). Furthermore,  $f(r|p)$  means a probability of success number is exactly  $r$  given an average probability  $p$ .

$$f(r|p) = \binom{N}{r} p^r (1 - p)^{N-r} \quad (3.8)$$

for  $r = 0, 1, \dots, N$  where

$$\binom{N}{r} = \frac{N!}{r! \times (N - r)!}$$

Coin flip is a simple example to illustrate the general idea of binomial distribution. A coin is tossed 10 times and the average probability of getting heads is 0.5 in a fair toss. The expected number of head is 5 but there still have a chance to get 3 heads only. According to Equation (3.8), the probability of getting 3 heads out of 10 tosses is  $f(3|0.5) = 0.1172$ .

### Beta-Binomial Inference

When the prior distribution follows beta distribution and likelihood being binomial distribution, the general Bayes theorem in Equation (3.6) is transferred into Equation (3.9).

$$f(p|r) = \frac{f(r|p) \times f(p)}{\int_0^1 f(r|p)f(p)dp} \quad (3.9)$$

$f(r|p)$  is a binomial distribution and  $f(p)$  is a beta distribution.  $N$  is the total number of sample set  $r$  is the number of success we observed and  $p$  is the probability of success outcome occurred.  $\int_0^1 f(r|p)f(p)dp$  is a normalizing factor that ensure the range of posterior within  $[0,1]$ .

$$f(p|r) = \frac{\Gamma(\alpha + \beta + N)}{\Gamma(\alpha + r)\Gamma(N - r + \beta)} p^{(r+\alpha-1)}(1 - p)^{(N-r+\beta-1)} \quad (3.10)$$

Equation (3.10) is the standard form of beta distribution with parameter  $\alpha' = \alpha + r, \beta' = \beta + N - r$ . :

$$f(p|r) = \phi_{\text{beta}}(\alpha', \beta') \quad (3.11)$$

From Equation (3.11), the biggest advantage of beta-binomial inference is that the posterior PDF can be obtained by counting the number of successes and total sample number without integration. When we have no idea about the probability of success,  $f(p)$ , we usually set the original parameters  $(\alpha, \beta) = (1, 1)$ , an uninformative uniform prior. By adding sample data, the reliability estimation will be continuously updated by counting number of success  $r$  and total sample number  $N$ .

### 3.2.2 Time variant life data

Assuming that the failure rate of a system is a constant over time, therefore, the time of the  $n$ th failure occurring can accurately be estimated. However, in reality, the true failure rate is not

known due to the lack of data. The failure rate is therefore some distribution and the time of the  $n$ th failure is also follows some types of distribution. In this thesis, we infer the probability of failure rate based on observations at different time instant using the Poisson-gamma inference.

The Poisson-gamma inference is frequently applied to infer failure rate within an interval of time according to the observed failure number. We will introduce Poisson distribution and gamma distribution separately and then derive the general form of posterior distribution based on Bayesian inference. The posterior distribution is the failure rate given the observations that can be used to estimate the time of the  $n$ th failure.

### Poisson Distribution

Poisson distribution is a discrete distribution frequently used to predict the probability of expected number of failure event occurring given a failure rate in a time or space interval. When each probability of event occurring is independent. The PDF is:

$$f(k; \lambda) = \frac{\lambda^k}{k!} \times e^{-\lambda} \quad \text{for } k = 1, 2, 3, \dots \quad (3.12)$$

where  $\lambda > 0$  is the average failure rate within a time interval and  $k$  is the expected number of failure events in the same interval.

### Gamma Distribution

Gamma distribution belongs to exponential distribution family with two parameters, a shape parameter  $a$  and a rate parameter  $b$ . It is commonly applied to model rainfall or waiting time. In this thesis, gamma distribution is use as prior of failure rate  $\lambda$ . The Gamma PDF is:

$$f(\lambda; a, b) = b^a \frac{1}{\Gamma(a)} \lambda^{a-1} e^{-b\lambda} \quad (3.13)$$

where  $\Gamma(\cdot)$  is the Gamma function. Note that when the shape parameter  $a = 1$ , gamma distribution will become an exponential distribution with a rate parameter  $b$ .

### Poisson-Gamma Inference

Poisson-Gamma inference is mainly used to infer the failure rate based on life data in this thesis. With the prior distribution  $f(\lambda)$  be a gamma distribution and the likelihood  $f(k|\lambda)$  be a Poisson distribution, the posterior is shown in Equation(3.14).

$$f(\lambda|k) = \frac{f(k|\lambda) \times f(\lambda)}{\int_0^\infty f(k|\lambda) \times f(\lambda)d\lambda} \quad (3.14)$$

Substitute PDF of Poisson and Gamma distributions into and Equation(3.14) results as in Equation(3.15):

$$f(\lambda|k) = \frac{\frac{e^{-\lambda}\lambda^k}{k!} \frac{b^a}{\Gamma(a)} \lambda^{a-1} e^{-\lambda b}}{\int_0^1 \frac{e^{-\lambda}\lambda^k}{k!} \frac{b^a}{\Gamma(a)} \lambda^{a-1} e^{-\lambda b} d\lambda} \quad (3.15)$$

where

$$\int_0^1 \frac{e^{-\lambda}\lambda^k}{k!} \frac{b^a}{\Gamma(a)} \lambda^{a-1} e^{-\lambda b} d\lambda = \frac{1}{k!} \frac{b^a}{\Gamma(a)} \frac{\Gamma(k+a)}{(b+1)^{k+a}} \quad (3.16)$$

Substitute Equation(3.16) into (3.15), we see that the posterior distribution become a Gamma distribution with new parameters of  $a' = a + k$  and  $b' = b + 1$ . When the interval of time is divided into  $n_g$  sections, the parameters is  $a' = a + \sum_{i=1}^{n_g} k_i$  and  $b' = b + n_g$ :

$$f(\lambda|k) = \phi_{\text{gamma}}(a', b') \quad (3.17)$$

$f(\lambda|k)$  is the density function of failure rate given observed failure number. By to Poisson counting process, if the number of failure is assumed to follow a Poisson distribution, the waiting time  $T[c]$  until the  $c$ th failure occurring is a Erlang distribution. The CDF of Erlang distribution,  $\Pr(T[c] \leq t)$ , is the probability of the time of the  $c$ th failure earlier than the time  $t$ , as shown in Equation(3.18):

$$\Pr(T[c] \leq t) = \Phi_{\text{Erlang}}(t; c, \lambda) = 1 - \sum_{n=0}^{c-1} \frac{1}{n!} (\lambda t)^n t^{n-1} e^{-\lambda t} \quad (3.18)$$

A more reliability product has less number of failures within the same time interval compared with its competitors. Therefore, the reliability of aging process,  $R_a$ , is the probability of the failure number is less than  $c$  before at time  $t$ .  $\Pr(T[c] \geq t)$ .

$$R_a = \Pr(T[c] \geq t) = 1 - \Phi_{\text{Erlang}}(t; c, \lambda) = \sum_{n=0}^{c-1} \frac{1}{n!} (\lambda t)^n t^{n-1} e^{-\lambda t} \quad (3.19)$$

where  $t$  is the time period we are interested in and  $c$  is the critical failure number. When the failure rate  $\lambda$  is a constant,  $\Pr(T[c] > t)$  is a fixed value. However, when the failure rate is a gamma distribution inferred from Poisson-Gamma inference in this thesis,  $\Pr(T[c] > t)$  is also a distribution. Figure 3.1 shows that we sample some  $\lambda$  from a Gamma distribution, reliability of aging process would vary within a range. The x-axis is the time, the y-axis is the reliability of aging process,  $R_a$ , and the each of line is portrayed with a fix  $\lambda$  value. When the  $\lambda$  are drawn

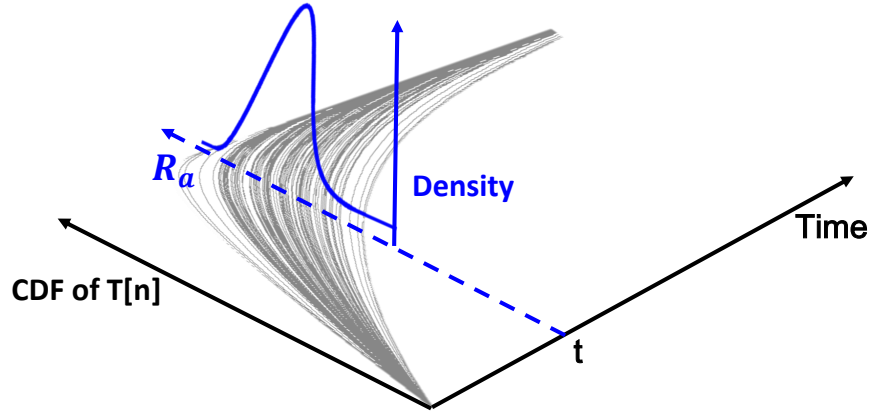


Figure 3.1: An illustration of the effect of parameter with distribution

from a gamma distribution and the sample size of  $\lambda$  is large enough, the PDF of reliability of aging process could be yielded.

In this thesis, beta-binomial inference is applied to estimate the reliability of an engineering design at the initial state. The Poisson -Gamma inference is used to estimate the failure rate of this function given the observed failure number. The failure rate is further used to estimate the failure probability within a period of time.

### 3.3 Reliability estimation using Bayesian inference with life data

Constraints in the engineering design contain some parameters with uncertainty. In reality, we can only get few measurements of the parameters. Therefore, these measurements need to be accounted for to estimate the reliability of the constraints by inferences which have been introduced in section 3.2. In this thesis, the basic inferring process will be introduced in subsection 3.3.1. Confidence range and confidence bound will be introduced in the inferring process in subsection 3.3.2. A mathematical example will be used to demonstrate the inferring process in section 3.4.

### 3.3.1 Reliability estimation of constraints

Let the number of time invariant samples  $P_1$  be  $N_0$ , of the time variant samples  $P_2(t)$  be  $N_2(t)$  at time  $t$ , and of the time variant samples  $P_3(t)$  is  $N_3(t)$  at time  $t$ . The time-independent samples,  $P_1$ , could combine with each sample at any time instant as they will not be changed over time. However, time-dependent samples can only combine with samples which are measured at the same time. Therefore, we can denote the sample sets we obtain at time  $t$  be  $(P_1, P_2(t), P_3(t))$  at time  $t$ . For example, the number of the time invariant samples is  $N_0 = 3$ , the time interval now is divided into two segments,  $t_1$  and  $t_2$ , the number of segments of time interval,  $N_g$ , is 2. The numbers of the time variant samples  $P_2$  are  $N_2(t_1) = 3$  and  $N_2(t_2) = 4$ , and the numbers of the time variant samples  $P_3$  are  $N_3(t_1) = 2$  and  $N_3(t_2) = 3$ . Therefore, the number of sample sets  $N_c(t)$  at  $t_1$  is  $N_c(t_1) = 3 \times 3 \times 2 = 18$ , and the number of sample sets at  $t_2$  is  $N_c(t_2) = 3 \times 4 \times 3 = 36$ .

Some samples would make the value of a constraint being less than zero, and some samples would make the value of a constraint being greater than zero. We define a constraint being less than zero a success event. Therefore, we can have the number of success based on the total number of samples at  $t$ .

Beta-binomial inference requires the number of success  $r$  and the number of sample set  $N_c(t_1)$  at initial time  $t_1$ . while Poisson-gamma inference requires the total number of failure at all time instant  $\sum_{i=1}^{N_g} k_i$  and the number of the time interval  $N_g$ . With these distribution parameters beta-binomial inference becomes:

$$f(\mathbf{R}|r) = \phi_{\text{beta}}(\alpha + r, N_c(t_1) - r + \beta) \quad (3.20)$$

and the Poisson-gamma inference becomes:

$$f(\lambda|k) = \phi_{\text{gamma}}(a + \sum_{i=1}^{N_g} k_i, b + N_g) \quad (3.21)$$

Equation(3.20), we can infer the distribution of reliability  $\mathbf{R}$  when observer  $r$  success and from



Equation(3.21), the distribution of the failure rate  $\lambda$  given current number of failure  $k$ . The

If the distribution of some design parameters are known, the beta-binomial inference and Poisson-gamma inference is modified slightly. In the previous case mentioned in the beginning of this subsection, tests of each sample set are defined as pass/fail with reliability being 1 or 0. Therefore, we can simply count the number of successful samples and use the value as number of success  $r$ .

Consider a performance function  $g$  with design parameters with both distribution,  $\mathbf{P}_u$  and with samples,  $\mathbf{P}_s$ , the reliability of the constant with the  $i$ th sample set for  $\mathbf{P}_s$  is:

$$R_i = \Pr[g(\mathbf{X}_u, \mathbf{P}_u) |_{(\mathbf{X}_s, \mathbf{P}_s)_i} \leq 0] \quad (3.22)$$

Figure 3.2 illustrate the case with three samples combined with distribution design parameter. The area on the left hand side of  $g = 0$  is the reliability value. As shown, that reliability value depend on the design parameter samples. The expected total number of success  $r$  is:

$$E[r] = \sum_{i=1}^N R_i \quad (3.23)$$

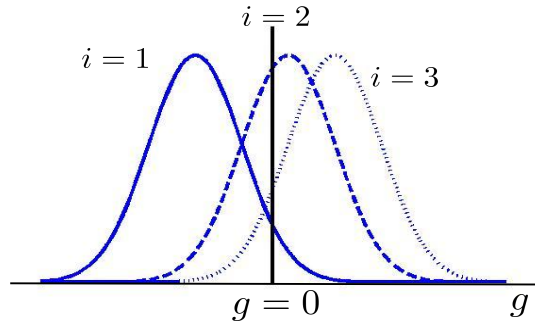


Figure 3.2: An image of reliability given distribution with  $i$ th sample data

The expression of beta-binomial inference with distribution and sample design parameters is:

$$\mathbf{R} \sim \phi_{\text{beta}}(E[r] + \alpha, N_1 - E[r] + \beta) \quad (3.24)$$

Similar idea can be applied to estimate the number of failure using Poisson-gamma inference. In Figure 3.2, the area on the right hand side of  $g = 0$  is the failure probability and sum up the failure probability given the design parameter samples, the expected failure number,  $E[k]$ ,

is yielded. The  $k$  in Equation (3.21) are replaced by  $E[k]$ .

$$\lambda \sim \phi_{\text{gamma}}\left(\sum_{i=1}^{n_g} E[k]_i + a, b + n_g\right) \quad (3.25)$$

By counting the exact or the expected failure number, the distribution of failure rate is yield. The failure rate can be used to estimate the reliability of an aging process,  $R_a$ .

$$R_a = \Pr(T[c] \geq t) = 1 - \Phi_{\text{Erlang}}(t; c, \lambda) = \sum_{n=0}^{c-1} \frac{1}{n!} (\lambda t)^n t^{n-1} e^{-\lambda t} \quad (3.26)$$

where  $t$  the is usually warranty we set, the  $c$  is the maximum acceptable number of failure. Furthermore, warranty and the max failure number are fixed value but the parameter  $\lambda$  is a distribution that we yield from Poisson-gamma inference. Since  $\lambda$  is a distribution, the reliability of aging process is also a distribution.

Follow the process mentioned in subsection 3.3.1, we can analyze a the reliability of a constraint with life data in both initial state and aging process. This analysis can be used in a design process so that we can control the quality of a design at initial state and ensure the failure number of failure within a time interval is less than specific number; it is a similar concept to warranty.

### 3.3.2 Definition of confidence range and confidence bound

Confidence range is the degree of confidence of the reliability from the inference with sample data. In beta-binomial inference, the distribution of reliability  $R$  is yielded from Equation (3.20). Setting a critical value  $R_t$  and the probability of  $R$  greater than  $R_t$ ,  $\Pr(R \geq R_t)$ , is called the confidence range of reliability,  $CR_R$ , shown in Equation(3.27):

$$CR_R = \Pr(R \geq R_t) = 1 - \Phi_{\text{beta}}(R_t, \alpha, \beta) \quad (3.27)$$

When assuming all the samples are successful events, we can get the maximum confidence range of reliability, is also named confidence bound of reliability,  $CB_R$ .

$$CB_R = \max[\Pr(R \geq R_t)] = 1 - \Phi_{\text{beta}}(R_t, N_1 + 1, 1) \quad (3.28)$$

According to Equation (3.28),  $CB_R$  is only associated with reliability target  $R_t$  and total number of samples  $N_1$ . If the confidence bound of reliability  $CB_R$  is less than confidence range target,  $CR_t$ , we assign, we have to increase the number of samples.

The similar concept can be applied to Poisson-gamma inference. The reliability of and aging process is denoted as  $\Pr(T[c] \geq t) = R_f$ , the reliability target of an aging process is symbolized as  $P_t$  and the probability of the reliability of aging process being larger than a probability target is defined as confidence range of an aging process,  $CR_f$ . The expression of  $CR_f$  is shown:

$$CR_f = \Pr[\Pr(T[c] \geq t) \geq P_t] = \Pr[R_a \geq P_t] \quad (3.29)$$

Assuming all current samples are safe, the failure rate inferred will be the lowest failure rate one can expect. This lowest failure rate can infer the probability of the  $c$ th failure occurring,  $T[n]$ , being later than the time we set of an aging process. The maximum confidence level is called the confidence bound of failure.

$$CB_f = \max[\Pr[R_a \geq P_t]] \quad (3.30)$$

$CB_f$  is the highest value of confidence level under current sample number. When confidence bound of failure  $CB_f$  is less than confidence level target  $CL_t$ , this target are impossible to reach with handing samples. Therefore, adding new samples is necessary.

### 3.4 Reliability estimation example

Taking a constraint  $G(P_1, P_2) = 1 - 80/(P_1^2 + 8 \times P_2 - 6)$  for demonstrating example where  $P_1$  is time invariant parameter,  $P_1 \sim N(-8.2, 0.08^2)$ , and  $P_2$  would deteriorate over time,  $P_2 \sim N(2.2 \times e^{-t}, 0.02^2)$ . Furthermore, the target probability are 0.7 for both reliability and failure,  $R_t = 0.7$  and  $P_f = 0.7$ . Then, calculating the confidence range of reliability that reliability distribution greater than  $R_t$ ,  $CR_R = \Pr[R > R_t]$ ; also calculating the confidence range of reliability of aging process,  $CR_f = \Pr[R_a > P_t]$ . There are two cases to verify how increments of sample size of time invariant and variant parameters will affect the confidence ranges of reliability in both initial state and aging process.

### Case 1

We originally have five samples of  $P_1$  and three samples of  $P_2$  at each time,  $t = \{0, 1, 2\}$ ;  $t = 0$  is the initial state, and then enlarge the sample size of  $P_1$  to ten samples and  $P_2$  remains the same as original sample. The sample in case 1 is shown in Table 3.1.

Since the initial state is the only focus in beta-binomial inference, the original number of

Table 3.1: Samples of  $\mathbf{P}_1$  and  $\mathbf{P}_2$  in case 1

$P_1$	Original	-8.082	-8.241	-8.206	-8.240	-8.292
	Additional	-8.204	-8.194	-8.016	-8.198	-8.208
$P_2$	t=0	2.165	2.222	2.1951		
	t=1	0.041	0.009	-0.014		
	t=2	0.018	-0.002	-0.016		

sample sets for beta-binomial inference,  $N_1 = 5 \times 3 = 15$  and 12 sets out of 15 sample sets are safe. Therefore, the confidence range of reliability can be obtained,  $CR_{R_{\text{scenario 1}}} = \Pr[R > R_t] = 1 - \Phi_{\text{beta}}(0.7, 12 + 1, 15 - 12 + 1) = 0.7541$  in scenario 1. After adding new sample, the number of sample set for beta-binomial inference is changed to  $10 \times 3 = 30$  and 27 sets out of 30 are safe. The updating confidence range of reliability is calculated,  $CR_{R_{\text{scenario 2}}} = \Pr[R > R_t] = 1 - \Phi_{\text{beta}}(0.7, 27 + 1, 30 - 27 + 1) = 0.9928$  in scenario 2. Figure 3.3 shows the confidence range of reliability increase with rising number of samples. The area in the right hand side is the confidence of reliability. As we can see, the confidence of reliability in scenario 2 is larger than which in scenario 1. Although adding new samples of  $P_1$  can also increase the number of sample set in Poisson-gamma inference, the confidence range of reliability of aging process,  $R_f$ , dose not change too much in this case,  $CR_{f_{\text{scenario 1}}} = \Pr[R_a > 0.7] = 0.0258$  and  $CR_{f_{\text{scenario 2}}} = \Pr[R_a > 0.7] = 0.0234$ . In fact, the confidence range of reliability of aging process is influenced by the failure rate. We can learn more about failure rate by adding samples at different instants instead of adding new samples of the time instant we have had samples already.

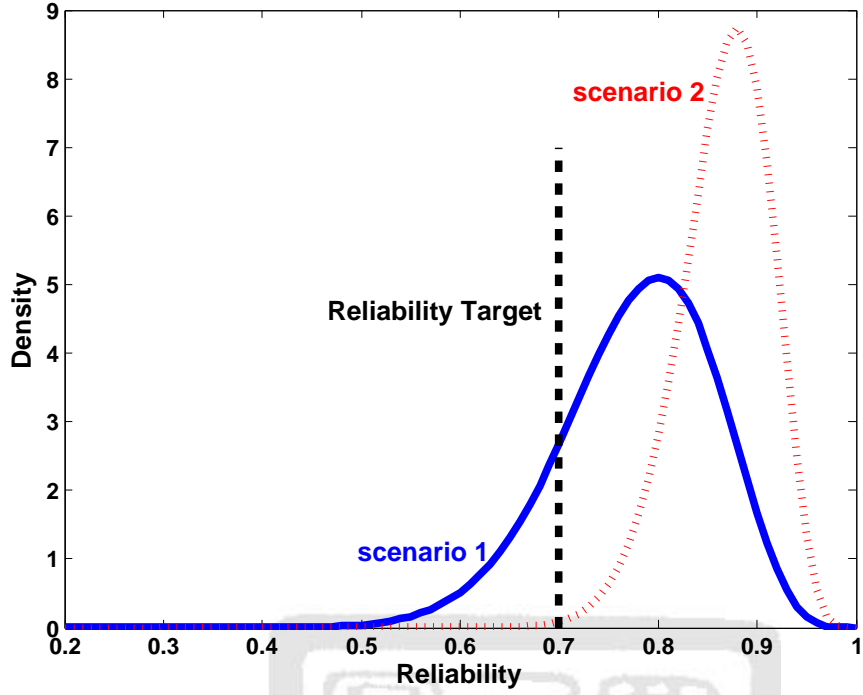


Figure 3.3: Reliability distribution with different sample set of reliability estimation

### Case 2

We have five samples of  $P_1$ , three samples of  $P_2$  at  $t = 0, 1, 2$ ; therefore, the number of time section is three,  $n_g = 3$  in scenario 1. Now, adding three new samples at  $t = 0.25, 0.5, 0.75, 1.25, 1.5, 1.75$ ; therefore, the number of time section is nine,  $n_g = 9$  in scenario 2. The sample data in case 2 is shown in Table3.2.

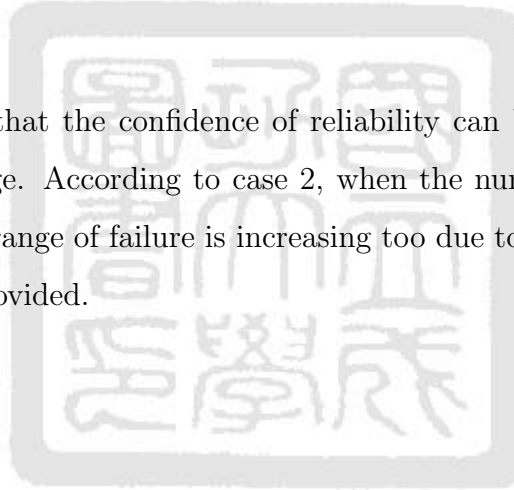
Table 3.2: Samples of  $\mathbf{P}_1$  and  $\mathbf{P}_2$  in case 2

$P_1$	$P_2$								
	t=0	t=0.25	t=0.5	t=0.75	t=1	t=1.25	t=1.5	t=1.75	t=2
-8.082	<b>2.165</b>	0.788	0.316	0.098	<b>0.040</b>	0.016	0.007	-0.023	<b>0.018</b>
-8.241	<b>2.222</b>	0.786	0.283	0.096	<b>0.009</b>	-0.028	0.030	0.028	<b>-0.002</b>
-8.206	<b>2.195</b>	0.844	0.314	0.118	<b>-0.014</b>	-0.002	0.024	0.009	<b>-0.016</b>
-8.240									
-8.293									

The total number of sample sets is  $15 \times 3 = 45$  and 3 sets out 45 sets fail in Poisson-gamma inference. The failure rate is gamma distribution with parameters  $a' = 1 + 3$  and

$b' = 1 + 3$ ; therefore, the mean of failure rate is  $a'/b' = 1$ . Using this failure rate which is a distribution, we can yield confidence range of reliability of aging process by MCMC algorithm,  $CR_{f\text{scenario1}} = \Pr[R_a > 0.7] = 0.0258$ . After that, adding additional tree samples at  $t = 0.25, 0.5, 0.75, 1.25, 1.5, 1.75$ , time interval is divided into 9 section,  $n_g = 9$ . The total number of sample sets is  $15 \times 9 = 135$  and 3 sets out of 135 fail. we can infer the new distribution of the failure rate; the new failure rate is a gamma distribution with parameters  $a' = 1 + 3$  and  $b' = 1 + 135$ , and the mean of the new failure rate is  $E[\lambda] = a'/b' = 4/136$ . Furthermore, The confidence range of failure is  $CR_{f\text{scenario2}} = \Pr[R_a > 0.7] = 0.310$ . The confidence range of reliability is the same while adding new data, because the confidence range of reliability is only influenced by the number of sample sets at initial stage; in case 2, the number of sample sets is the same in scenario1 and scenario2.

From case 1, we can learn that the confidence of reliability can be improved by adding new sample points at initial stage. According to case 2, when the number of time interval is divided increase, the confidence range of failure is increasing too due to the increasing amount of knowledge of failure rate is provided.



# Chapter 4 Proposed Bayesian Updating Scheme with Life Data

We have introduced how to estimate the reliability and the confidence of the reliability. In this chapter, we propose an optimization scheme to address time-dependent design problem with life data and apply the estimated reliability to the optimization model. The overall design flowchart is presented in section 4.1. The details of the flowchart are discussed as follows: the optimization model is demonstrated in section 4.2, activity of constraints in optimization model is defined in section 4.3; When the confidence bounds or ranges in the optimization model do not reach the targets, we have to add new samples to increase the confidence ranges. The strategy of adding new samples is presented in section 4.4.

## 4.1 Overall design flowchart

Figure 4.1 shows the overall flowchart of the proposed design method. The reliability targets at the initial state,  $R_t$ , the reliability of an aging process,  $P_t$ , the confidence range targets at the initial state,  $CR_t$ , and the confidence range of an aging process,  $CL_t$  are pre-determined. Setting appropriate reliability and confidence values require strategic product planning across quality, cost and many other attributes. Once we have all the reliability and confidence requirements, the current uncertainty, including life, data are then classified into time-dependent and time-independent data. Based on the quantity of data size, we can further classify into time-independent distribution uncertainty,  $\mathbf{P}_u$ , time-independent uncertainty with samples,  $\mathbf{P}_s$ , time-dependent distribution uncertainty,  $\mathbf{P}_u(t)$ , and time-dependent uncertainty with samples,  $\mathbf{P}_s(t)$ .

The size of uncertainty determines its bound in calculating confidence. We then have the confidence bounds at the initial state,  $CB_R$ , and the confidence bounds of an aging process  $CB_f$ . When the data size is large enough to provide confidence bounds greater than the desired targets, an optimization process is then performed to obtain the best design within a given time frame based on the life data available. However, if the data size is inadequate, more samples are needed. The critical constraint with the lowest confidence range is determined and new

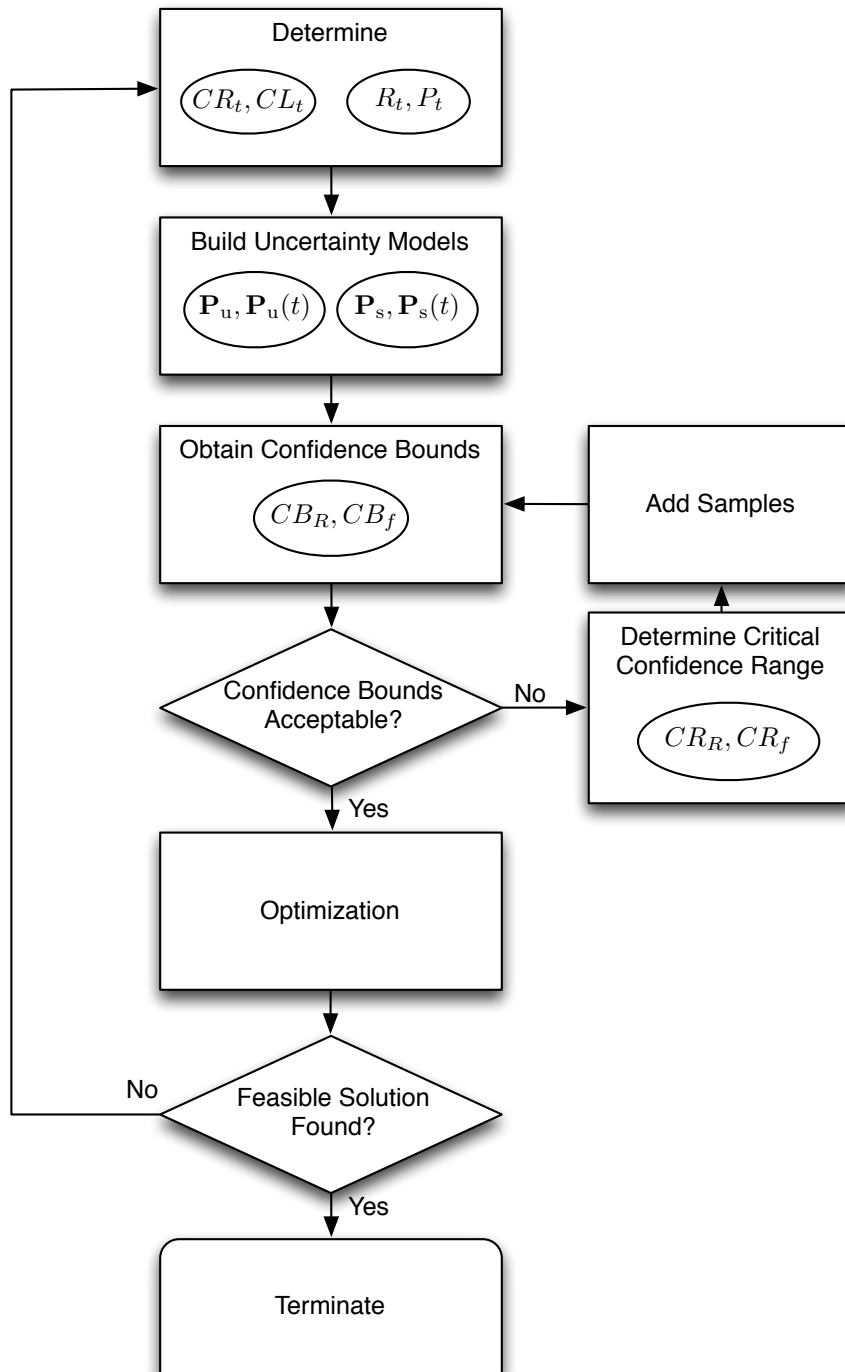


Figure 4.1: The overall design flowchart

samples of important parameters are added until the confidence bounds satisfy the confidence targets. This process ends when we have feasible solution to the optimization model. If there is no feasible solution found within the constrained space, we should refine the reliability targets and confidence range targets before redo the entire process.



The detail of the optimization model is presented in section 4.2, in which we also define the activity of constraints with confidence range in section 4.3. The strategy of adding samples including sensitivity analysis and MCMC bias sample filter is then discussed in the section 4.4

## 4.2 Optimization model

The general optimization model of time-dependent design problem is shown in Equation(4.1):

$$\begin{aligned}
& \min_{\boldsymbol{\mu}_{\mathbf{D}_u}, \mathbf{d}} \max(f(\boldsymbol{\mu}_{\mathbf{D}_u}, \mathbf{d}, \mathbf{P}_s, \boldsymbol{\mu}_{\mathbf{P}_u}, \mathbf{p})) \\
& \text{s.t } g_i(\boldsymbol{\mu}_{\mathbf{D}_u}, \mathbf{d}, \mathbf{p}) \leq 0 \\
& g_R = \Pr[g_i(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}) \leq 0] \geq R_t \\
& g_B = \Pr [\Pr[g_i(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{P}_s, \mathbf{p}) \leq 0] \geq R_t] \geq CR_t \\
& g_i(t) = g_i(\boldsymbol{\mu}_{\mathbf{D}_u}, \mathbf{d}, \mathbf{p}, \mathbf{p}(t)) \leq 0 \\
& g_R(t) = R_a(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}, \mathbf{P}_u(t), \mathbf{p}(t)) \geq P_t(t) \\
& g_B(t) = \Pr [R_a(\mathbf{D}_u, \mathbf{d}, \mathbf{P}_u, \mathbf{p}, \mathbf{P}_s, \mathbf{P}_u(t), \mathbf{p}(t), \mathbf{P}_s(t)) \geq P_t(t)] \geq CL_t(t)
\end{aligned} \tag{4.1}$$

where

$\mathbf{d}$  : deterministic design variables.

$\mathbf{D}_u$  : uncertain design variables with known distributions.

$\mathbf{D}_u(t)$  : time-dependent uncertain design variables with known distributions.

$\mathbf{p}$  : deterministic parameters.

$\mathbf{p}(t)$  : time-dependent deterministic parameters.

$\mathbf{P}_u$  : uncertain parameters with known distributions.

$\mathbf{P}_u(t)$  : time-dependent uncertain parameters with known distributions.

$\mathbf{P}_s$  : uncertain parameters with samples.

$\mathbf{P}_s(t)$  : time-dependent uncertain parameters with samples.

$g$  : deterministic constraint.

$g_R$  : reliability constraints with a reliability target  $R_t$ .

$g_B$  : Bayesian reliability constraints with a reliability target  $R_t$  and a confidence range target  $CR_t$ .

$g(t)$  : time-dependent deterministic constraint.

$g_R(t)$  : time-dependent reliability constraints with a reliability target  $P_t$ .

$g_B(t)$  : time-dependent Bayesian reliability constraints with a reliability target  $P_t$  and a confidence range target  $CL_t$ .

$R_a$  : the reliability of aging process which is inferred based  $\mathbf{p}$ ,  $\mathbf{p}(t)$ ,  $\mathbf{P}_u$ ,  $\mathbf{P}_u(t)$ ,  $\mathbf{P}_s$  and  $\mathbf{P}_s(t)$ .

$R_t$  and  $P_t$  : the reliability target for time-independent and time-dependent cases.

$CR_t$  and  $CL_t$  : the confidence range target for time-dependent and time-independent cases.

The constraints  $g_R$ ,  $g_R(t)$ ,  $g_B$  and  $g_B(t)$  deal with different types of parameters. The original constraints are classified into one with time-dependent parameters and another one with time-independent parameters. The reliability of a function with time-dependent parameters of known distributions are transformed into a time-independent reliability measure at the initial state and a time-dependent reliability measure at a given life length  $t$ . The reliability constraint restrains the reliability in the initial state, and the time-dependent reliability constraints ensure the reliability target is satisfied of aging process. However, when the time-dependent constraint with parameters which are presented in sample points instead of the distributions, the original constraint would lead two Bayesian reliability constraints. These two Bayesian constraints guarantee the reliability target and the confidence range target be satisfied in both initial state and aging process. The time-independent Bayesian reliability constraint is constructed with beta-binomial inference, and the time-dependent Bayesian reliability constraint is constructed with Poisson-gamma inference.

Take three constraints for example, the original constraints in deterministic optimization are shown in Equation(4.2), where the parameters ,  $P_1$  and  $P_2$  are constants:

$$\begin{aligned} g_1 &= -5(d \times P_1 - 50) \\ g_2 &= 1 - (d - P_1 - 5)^3 - 50 \times P_2 \\ g_3 &= 50(d - P_2 - 2 \times P_1) \end{aligned} \quad (4.2)$$

If the parameters,  $P_1$  and  $P_2$ , are not constants anymore where  $P_1 \sim N(10, 0.8^2)$  and the underlying distribution of  $P_2 \sim N(5e^{-t}, 0.05^2)$ ,  $t \in \{0, 2\}$ , but we have only few sample points of both  $P_1$  and  $P_2(t)$  drawn from the underlying distribution. The original deterministic optimization is changed to be a Bayesian reliability optimization; therefore, changing the expression of these constraints is necessary.

For the constraint  $g_1$  with time-independent parameters with sample only, it would transfer into a time-independent Bayesian reliability constraint,  $g_{B1}$ . For the constraint  $g_2$  with aging sample parameter, it would lead to time-dependent and time-independent Bayesian reliability constraints,  $g_{B2}$  and  $g_{B2}(t)$ . The third constraint  $g_3$  is in the same situation as the constraint  $g_2$ . The constraint  $g_3$  is changed to  $g_{B3}$  and  $g_{B3}(t)$ . The final constraints are resulted in Equation(4.3):

$$\begin{aligned} g_{B1} &= \Pr [\Pr [-5(d \times P_1 - 50)] \geq R_t] \geq CR_t \\ g_{B2} &= \Pr [\Pr [1 - (d - P_1 - 5)^3 - 50 \times P_2] \geq R_t] \geq CR_t \\ g_{B3} &= \Pr [\Pr [50(d - P_2 - 2 \times P_1)] \geq R_t] \geq CR_t \\ g_{B2}(t) &= \Pr [R_{a2} \geq P_t] \geq CL_t \\ g_{B3}(t) &= \Pr [R_{a3} \geq P_t] \geq CL_t \end{aligned} \quad (4.3)$$

Depending on the types of data at hand, the constraints in deterministic case are changed to different forms. The optimization formulation in Equation(4.1) discussed in this thesis provide the most general case with all possible data forms be included.

The definition of the objective function should also be modified due to different data forms. When the discrete samples is substitute to the objective function, there are more than one objective function value. Due to the objective function has to return to a value in the optimization, we chose the maximum of the objective function values to be the objective. We can ensure the all possible objective function will be less than this value.

### 4.3 Activity of Bayesian aging constraints

The activity of the constraints in a design problem enables designers to understand which function performance restrict the design toward a better value. Theoretically, a constraint is considered as active if its removal will alter the optimal solutions. In most deterministic cases, active constraints are usually the ones satisfied as a strict equality at the optimum, as in Equation (4.4). Although this definition is not theoretically rigorous, it is quite practical in most design problems.

$$g(\mathbf{x}^*) = 0 \quad (4.4)$$

Considering the design parameters of a constraint with explicit distributions, the certain shape of the constraint is yielded by substitute the distribution of design parameters into the constraint, and the deterministic optimal problem is transferred into an RBDO problem. In RBDO, the reliability value  $R$  of a constraint is defined as the area on the left hand side of the criteria  $g = 0$  shown in the Figure 4.2. The reliability,  $R$ , of the constraint has to greater than

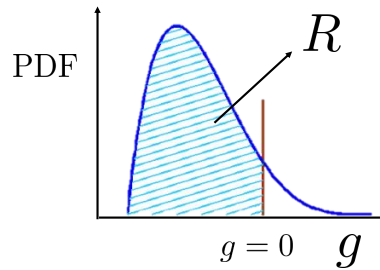


Figure 4.2: The reliability of a constraint

the reliability target,  $R_t$ . In other words, the  $(1 - R_t)$  percentile of the constraint is greater than zero. The activity of the constraint is define as the reliability just satisfy the requirement shown in Equation (4.5).

$$g^{(1-R_t)\%} = 0 \quad (4.5)$$

When the design parameters include discrete samples, the distribution of a constraint will alter depending on the value of the sample as well as the sample size. A constraint with two

samples will result in two different probability density functions. Therefore, the reliability of a constraint is not a constant anymore; a distribution instead. The original RBDO formulation is then transferred into Bayesian RBDO that has the reliability as a distribution  $\mathbf{R}$ . The form of the constraint in Bayesian RBDO is the probability of the reliability,  $R$ , of the constraint which is greater than target,  $R_t$ , is greater than a confidence range target,  $CR_t$ . Therefore, we can apply the same logic to calculate the reliability of the reliability, and set the activity of the constraints as Equation(4.6):

$$R^{(1-CR_t)\%} = R_t \quad (4.6)$$

However, in the beta-binomial inference with sample points, the parameters of inferred reliability distribution are discrete. Consequently, the area on the right hand side of reliability target, called the confidence range, is also discrete. The same phenomenon is happened in Poisson-gamma inference. The distribution parameters of the failure rate which is inferred from Poisson-gamma inference are discontinuous. Due to the inferred distribution of the failure rate, the confidence range of aging process is also discrete. Therefore, a specific value of confidence range target may not be able to be reached in both initial state and aging process. Beside, the confidence ranges of the aging process do not have a closed form, we infer the confidence ranges with MCMC that is giving a lot of samples to estimate the confidence ranges. Therefore, the confidence ranges of aging process might not so accuracy do to the MCMC algorithm. Considering these two reasons, we have to release the definition of activity. The new definition of activity shows in Equation(4.7):

$$\min(\mathbf{R}^{(1-CR)\%} - R_t) \quad (4.7)$$

There are two modified in Equation (4.7). First, Using the  $(1 - CR)$  percentile  $\mathbf{R}$  instead of applying  $(1 - CR_t)$ . The confidence ranges represent the degree of trust of these data, once the confidence ranges satisfy the confidence target, they do not affect the following design process; they serve as the thresholds. Furthermore, the specific confidence range target is not available, the real confidence range,  $CR$ , is adapted to represent the truth. Then, in the general case, the optimal design is right on the activity constraint. Taking the error form MCMC into consideration, we release the definition of the activity; we just find the constraint which is closest to optimal design.

## 4.4 Resource allocation of sample augmentation

When the confidence bounds do not exceed confidence target, adding new sample is needed. Due to measuring takes time and cost, we want to just adding new samples on the most important parameter. Therefore, before adding new sample, we should determine which parameter is the most important to us. The strategy of adding new samples on critical parameter is shown in the Figure 4.3

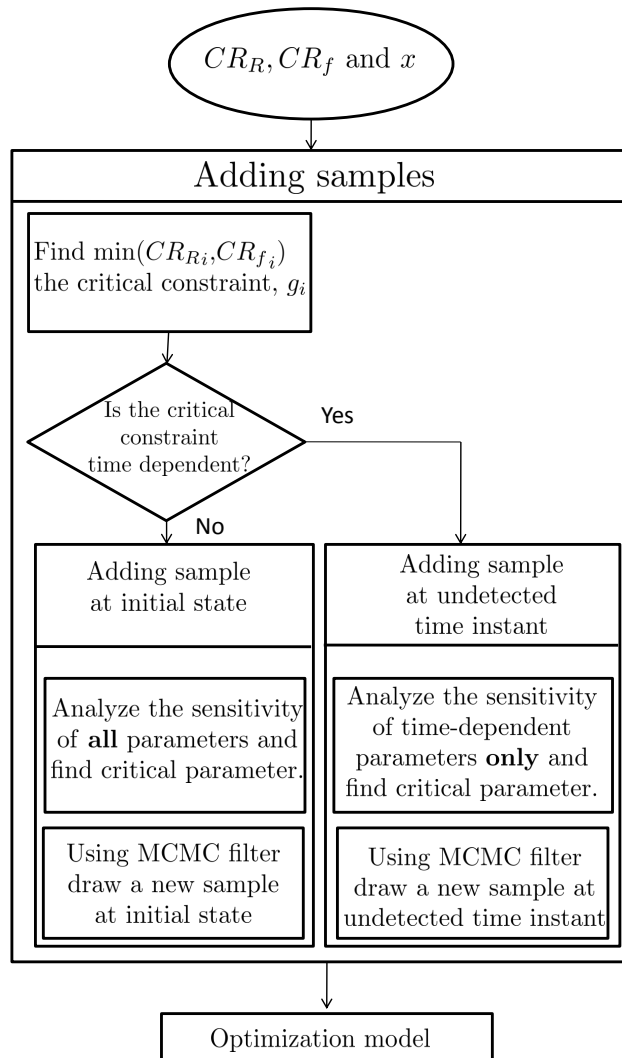


Figure 4.3: The strategy of adding new samples

After the first time optimization, we can yield the confidence ranges,  $CR_R$  and  $CR_f$ , in the initial state and aging process. The constraint which has the lowest confidence ranges is

the critical constraint. If the critical constraint is the time-invariant, analyzing the sensitivity on all sample parameters and set the most sensitive parameter as the critical parameter and then adding new sample on the critical parameter in the initial state. However, if the critical constraint is time variant, in order to get more information about failure rate, adding new sample in the undetected time instant is required, hence, only the time-dependent parameters are focused. The detail about sensitivity analysis and the algorithm of adding samples are demonstrated in subsection 4.4.1 and 4.4.2.

**Step 1** Define the constraint with the lowest confidence range as the critical constraint.

**Step 2** Determine whether the critical constraint is a time-invariant constraint or not. If the critical constraint is time-invariant, go to the step 3; otherwise, go to step 5.

**Step 3** Find the critical sample parameter on the critical constraint by sensitivity analysis. If the the critical constraint is the same as the previous iteration, we can simply adding new samples on the critical parameter we yield in last iteration without executing sensitivity analysis.

**Step 4** Adding new samples on the critical parameter in the initial state with Markov chain Monte Carlo (MCMC) Metropolis-Hastings algorithm. Go to the step 7.

**Step 5** When the critical constraint is a time-variant constraint, we execute sensitivity analysis on time-dependent parameters only. Find the critical time-dependent sample parameters.

**Step 6** Adding new samples on the critical time-dependent parameter at undetected time instant with Markov chain Monte Carlo (MCMC) Metropolis-Hastings algorithm. Adding one sample on the critical time-dependent parameter means the time segments plus one.

**Step 7** After adding new sample, return to update the optimization model to calculate the updating confidence bounds and the confidence ranges. If the updating confidence bounds still do not reach the confidence target, continuously adding new sample with this strategy.

### 4.4.1 Sensitivity analysis

In this thesis, we define the constraint with the lowest confidence range as the critical in Bayesian optimization model. In order to rise the confidence range, adding new sample is necessary. Before augment the sample size, We have to decide which parameter with sample is the critical parameter. The critical parameter means the variation of the parameter would lead to the most severe change on the constraint. We can decide the critical parameter by analyzing sensitivity respect to each parameter in the critical constraint. The sensitivity of  $g$  with respect to each parameter value is the derivative with respect to each parameter. Definition of the sensitivity at the value of  $(\bar{y}, \bar{z})$  shown in Equation(4.8):

$$\begin{aligned} s_y &= \left| \frac{\partial g}{\partial y} \right|_{y=\bar{y}} \\ s_z &= \left| \frac{\partial g}{\partial z} \right|_{z=\bar{z}} \end{aligned} \tag{4.8}$$

where  $s_y$  and  $s_z$  are defined as sensitivity of parameter  $y$  and  $z$ .

Note that if the critical constraint is time variant, we analyze sensitivity of the critical constraint with respect all the parameters in sample. On the other hand, when the critical constraint is time variant, we just analyze the sensitivity on the time-dependent parameters only.

### 4.4.2 MCMC bias sample filter

The estimation with bias samples might lead to inaccurate results; therefore, we have to use Markov Chain Monte Carol (MCMC) (the Metropolis-Hastings algorithm is adopted in this thesis) to avoid bias sample. MCMC is the method to generate a sequence of random samples from a distribution and the Metropolis-Hastings algorithm is used to estimate the movement tendency of the random sample within the filed of the distribution by calculating the acceptance index of the next sample candidate  $x_j$ . While calculating the acceptance index, the target distribution and proposal distribution must be known. We obtain the target distribution by use the statistic toolbox in commercial tool Matlab of existing samples, and the proposal distribution is obtained by re-sampling technique of Bootstrap method. The Bootstrap method



is used to estimate a unknown underlying distribution with few current samples. In the initial state, the size of current sample is  $n$ , and we can re-sample the current samples in the size  $n$ . When the sets of re-sample set is large enough, we can use them to estimate the parameters of underlying distribution.

The sequence of states constitutes a Markov chain with transition probability  $P_{i,j}$  as

$$\begin{aligned} P_{i,j} &= q(i,j)a_k(i,j), \text{ if } j \neq i \\ P_{i,i} &= q(i,i) + \sum_{k \neq i} q(i,k)(1 - a_k(i,k)) \end{aligned} \quad (4.9)$$

where  $q(i,j)$  is referred to as the *proposal* or *candidate-generating distribution*, represents that when a process is at the point  $i$ , the density generates a value  $y$  from  $q(i,j)$ . The Markov chain  $X_n$  will be time reversible and have stationary probabilities  $\pi(j)$  if

$$\begin{aligned} \pi(i)P_{i,j} &= \pi(j)P_{j,i} \text{ if } j \neq i \\ \pi(i)q(i,j)a_k(i,j) &= \pi(j)q(j,i)a_k(j,i) \end{aligned} \quad (4.10)$$

The  $a_k(i,j)$ , which is referred as the probability of move from  $x_i$  to  $x_j$ ,

$$a_k(i,j) = \min \left( \frac{\pi(j)q(j,i)}{\pi(i)q(i,j)}, 1 \right) \quad (4.11)$$

If the acceptance index is greater than one,  $a_k(i,j) \geq 1$ ; that means the probability of moving from  $x_i$  to  $x_j$  is higher, and then we can set the  $x_j$  as a new sample. In original concept of Metropolis-Hasting, When the candidate  $x_j$  of new sample is rejected,  $x_i$  would be added again.

In the case of the critical constraint being time variant, getting new samples at an undetected time instance is desired. Adding a new sample that is the same as current sample cannot help us learn more about the failure rate. Therefore, we have to draw an additional sample with MCMC rejected samples until there exist an acceptable sample. Figure 4.4 shows the flowchart of sample addition.

The procedure of the Metropolis-Hasting algorithm with arbitrary value  $s_0$  and our purpose is to draw samples from target distribution :

1. Estimate statistical parameters of the underlying distribution by directly using existing samples as target distribution  $\pi$ .

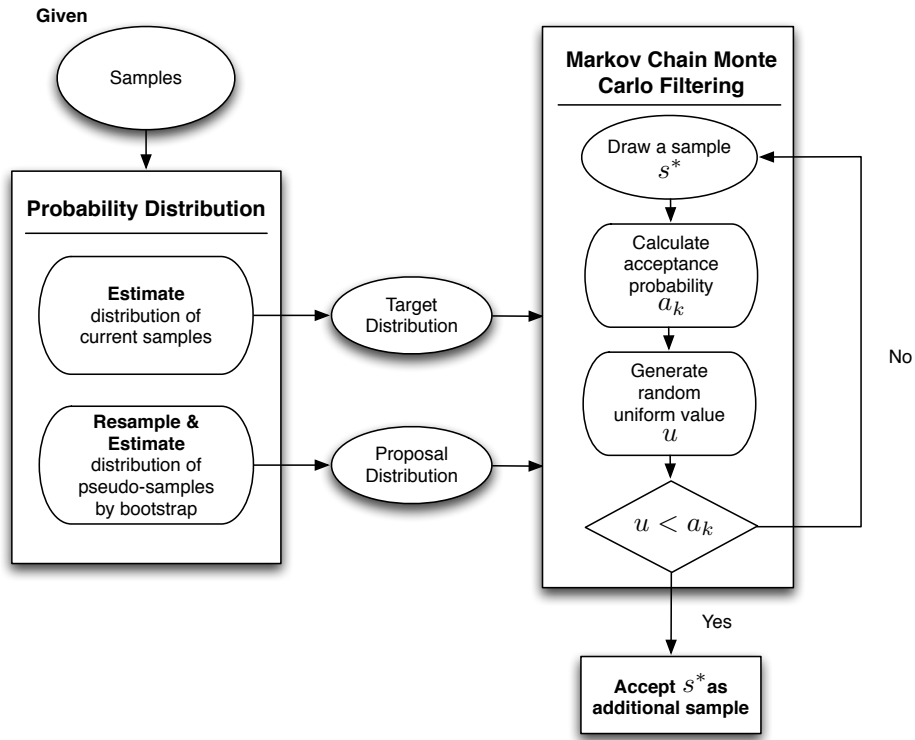


Figure 4.4: The procedure of MCMC of accepting an sample

2. Re-sample and estimate the statistical parameters by bootstrap method as the proposal distribution  $q$ .
3. Draw a sample  $s^*$  from population, calculate the acceptance probability  $a_k$  via the current sample  $s_c$ .

$$a_k(s_c, s^*) = \min \left( \frac{\pi(s^*)q(s^*, s_c)}{\pi(s_c)q(s_c, s^*)}, 1 \right)$$

4. Generate  $u$  from  $U(0, 1)$ .
5. If  $u \leq a_k(s_c, s^*)$ , then accept sample  $s^*$  as additional sample.
6. Else, turn to step 3.

In this thesis, we have time-independent and time-dependent samples so there two ways to add new samples. The procedure is the same in adding two types of samples, but the

ways to generate the target distribution and the proposal distribution are different. For time-independent case, we just use the MATLAB toolbox fitting the current data in normal distributions and yield the target distribution, and use the concept of the Bootstrap to estimate the proposal distribution. On the contrary, for the time-dependent case, we first the fitting the samples that measured in the same time period, if we have three segments within a time interval and we can yield three distributions, then get the modes of these three distributions, finally, use the three modes to fit the target distribution of failure rate with MATLAB toolbox . The way to generate the proposal distributions is much the same, but in the last step, we re-sample the modes to infer the distribution of the failure rate with Bootstrap.



## Chapter 5 Case Study

In this chapter the proposed design procedure is applied to solve a mathematical design optimization problem and an engineering problem with belt-pulley-tensioner system. A mathematical example is demonstrated and the proposed design methods are compared in section 5.1. In section 5.2, we use the sample procedure to evaluate a given belt-pulley-tensioner system with different data types considering deterioration.

### 5.1 A mathematical example

#### 5.1.1 Optimization models of mathematical example

##### Deterministic optimization model

$$\begin{aligned}
 \min_x f &= -20 \times (x - 15) \\
 \text{s.t. } g_1 &= -5(x \times p_1 - 50) \leq 0 \\
 g_2 &= 1 - (x - p_1 - 5)^3 - 50 \times p_2 \leq 0 \\
 g_3 &= 50(x - p_2 - 2 \times p_1) \leq 0
 \end{aligned} \tag{5.1}$$

where parameters  $p_1 = 10$  and  $p_2 = 5e^{-t}, t = 0$ .

##### RBDO model with abundant data

$$\begin{aligned}
 \min_x f &= -20 \times (x - 15) \\
 \text{s.t. } g_1 &= \Pr(-5(x \times \mathbf{P}_1 - 50) \leq 0) \geq R_t \\
 g_2 &= \Pr(1 - (x - \mathbf{P}_1 - 5)^3 - 50 \times \mathbf{P}_2 \leq 0) \geq R_t \\
 g_3 &= \Pr(50(x - \mathbf{P}_2 - 2 \times \mathbf{P}_1) \leq 0) \geq R_t \\
 \text{where } &\begin{cases} \mathbf{P}_1 \sim N(10, 0.08^2) \\ \mathbf{P}_2 \sim N(5e^{-t}, 0.05^2), t = 0 \end{cases}
 \end{aligned} \tag{5.2}$$

The reliability target is given as  $R_t = 0.9$ .

### Bayesian RBDO model with inadequate data

$$\begin{aligned}
 & \min_x f = -20 \times (x - 15) \\
 \text{s.t. } & g_1 = \Pr[\Pr[-5(x \times \mathbf{P}_1 - 50) \leq 0] \geq R_t] \geq CR_t \\
 & g_2 = \Pr[\Pr[1 - (x - \mathbf{P}_1 - 5)^3 - 50 \times \mathbf{P}_2 \leq 0] \geq R_t] \geq CR_t \\
 & g_3 = \Pr[\Pr[50(x - \mathbf{P}_2 - 2 \times \mathbf{P}_1) \leq 0] \geq R_t] \geq CR_t
 \end{aligned} \tag{5.3}$$

The confidence range target is given to be  $CR_t = 0.99$  and the reliability target is given as  $R_t = 0.9$  in this case. The initial number of samples of each uncertainty data is five. The initial samples are given in Table 5.1.

Table 5.1: 10 available initial data of  $\mathbf{P}_1$  and  $\mathbf{P}_2$  in Equation (5.3)

$\mathbf{P}_1$	$\mathbf{P}_2$
10.0904503643930	4.96629747941365
10.0976156668173	4.93917350485981
9.92732444983415	4.95235028759231
9.94977973640038	4.95077422452886
10.0221900268462	5.07280252345102

### Bayesian RBDO model with life data

$$\begin{aligned}
 & \min_x f = -20 \times (x - 15) \\
 \text{s.t. } & g_1 = \Pr[\Pr[-5(x \times \mathbf{P}_1 - 50) \leq 0] \geq R_t] \geq CR_t \\
 & g_2 = \Pr[\Pr[1 - (x - \mathbf{P}_1 - 5)^3 - 50 \times \mathbf{P}_2 \leq 0] \geq R_t] \geq CR_t \\
 & g_3 = \Pr[\Pr[50(x - \mathbf{P}_2 - 2 \times \mathbf{P}_1) \leq 0] \geq R_t] \geq CR_t \\
 & g_4 = \Pr[\Pr[T[n_{g_2}] > t] \geq P_t] \geq CL_t \\
 & g_5 = \Pr[\Pr[T[n_{g_3}] > t] \geq P_t] \geq CL_t
 \end{aligned} \tag{5.4}$$

The confidence range target at initial state is given to be  $CR_t = 0.9$ , the reliability target is given as  $R_t = 0.9$ , the confidence range target of aging process is given to be  $CL_t = 0.9$ , and the reliability target of aging process is given as  $P_t = 0.9$  in this case. The symbol,  $T[n_{g_i}]$ , means the time of  $n$ th failure of  $g_i$  occurring, and  $t$  is the warranty, here set  $n_{g_i} = 5$  and  $t = 10$ ; the constraints  $g_4$  and  $g_5$  control the failure rate of  $g_2$  and  $g_3$ . The initial number of samples of parameter  $P_1$  data is five and of parameter  $P_2$  is five at each observing time instant. The initial samples are given in Table 5.2. The observing time interval is divided into three segments,  $N_g = 3$ , and  $t_1 = 0$  is the initial state.

Table 5.2: 10 available initial data of  $\mathbf{P}_1$  and  $\mathbf{P}_2$  in Equation (5.4)

$\mathbf{P}_1$	$\mathbf{P}_2(t = 0)$	$\mathbf{P}_2(t = 1)$	$\mathbf{P}_2(t = 2)$
10.0904503643930	4.96629747941365	1.85730226976286	0.654998664348164
10.0976156668173	4.93917350485981	1.83792859681680	0.714295404951154
9.92732444983415	4.95235028759231	1.81662297116417	0.683691751839027
9.94977973640038	4.95077422452886	1.80236356335981	0.731012467314380
10.0221900268462	5.07280252345102	1.86061889615533	0.662052501344433

### 5.1.2 Comparison of results and discussion

In these optimization models, we set  $t = 0$  and  $t = 2$  to verify the influence of the deterioration of the aging design parameter,  $P_2$  with the underlying aging model. However, the proposed method infers the change of the reliability of constraints with the aging design parameter without verifying the real aging model and estimate the probability of the failure number being less than  $n$ .

In the deterministic optimization model, the optimal design is the most optimistic solution,  $x^* = 25$ , and the minimum objective function value is the smallest,  $f = -200$ , in the initial state,  $t = 0$ . Considering the uncertainty with a distribution, the solution yielded from the optimization is more conservative than deterministic optimization,  $x^* = 24.7268$  and  $f = -195.71$ . However, the result of the Bayesian RBDO is approximately like the result of RBDO,  $x^* = 24.7789$  and  $f = -195.57$ . In general case, the result of the Bayesian RBDO would be

more conservative than of RBDO, but the result could be influence by the quality of samples. Therefore, the result of Bayesian RBDO has a chance being more optimistic than RBDO. Taking deterioration into consideration, the Bayesian RBDO is transferred into time-dependent Bayesian RBDO. In order to control the failure number being less than 5 while  $t \in [0, 2]$ , the result of the time-dependent Bayesian RBDO is dramatically shifted,  $x^* = 20.4713$  and  $f = -109.4257$ . In fact, in this mathematical example, the underlying aging process is the exponential decay. We set the time,  $t = 2$ , the aging parameter affects the performance of these constraint, therefore, the optimal design of deterministic is shift,  $x^* = 20.676$  and  $f = -113.533$ . The optimal design of RBDO at  $t = 2$  is yielded,  $x^* = 20.4017$  and  $f = -109.270$ , and the optimal design of Bayesian RBDO at  $t = 2$  is obtained,  $x^* = 20.4595$  and  $f = -110.8951$ .

The assumption is slightly difference between time-dependent Bayesian RBDO and others. The requirements of time-dependent Bayesian RBDO is the number of failure being less than 5 within  $t \in [0, 2]$ . However, the requirement of RBDO and Bayesian RBDO is the probability of failure at  $t = 2$  is less than  $R_t$ . This two requirements would be transferred to another, the product of the size of the the population and the probability of failure is the number of failure. By assuming the size of the population is a hundred, and the limited number of failure is five, therefore, the failure probability is  $R = 0.95$ . The target reliability is set as  $R_t = 0.95$  in the RBDO and Bayesian RBDO optimization at  $t = 2$ .

Compared with the results at  $t = 2$  the result of the time-dependent Bayesian RBDO, the result of the results of RBDO, Bayesian RBDO, and time-dependent Bayesian RBDO are quit similar. We can use these aging data to predict the aging process without inferring the true aging model and yield a acceptable result.

### 5.1.3 Summary

This proposed method could infer the aging process without inferring the true model of aging process. The result yielded from the time-dependent Bayesian RBDO is not the conservative one, because the assumption is slightly different and the inferring process involves the random process. However, the time-dependent Bayesian RBDO could deal with the aging process and yield a reasonable result.

Table 5.3: The comparison of deterministic, RBDO, Bayesian RBDO, and Bayesian RBDO with life data of mathematical example

	Deterministic	RBDO	Bayesian RBDO	Bayesian RBDO with life data
$x_{\text{opt}}(t = 0)$	25	24.7268	24.7789	$x_{\text{opt}}=20.4713$ $f = -109.4257$
$f(t = 0)$	-200	-195.710	-195.5781	
$x_{\text{opt}}(t = 2)$	20.676	20.4017	20.4595	
$f(t = 2)$	-113.533	-109.270	-110.895	

## 5.2 A position of tensioner in belt-pulley system design optimization

We have been applied this optimization model to the mathematical example, now we want to extend it to an engineering belt-pulley system design. The belt-pulley system usually equip a tensioner to maintain a sufficient tension level to ensure the transmissive efficiency. Finding the proper position of the tensioner is an important design issue. In the section 2.1, the governing equations of this belt-pulley-tensioner system have been derived. We will use these governing equations to analyze the performance of a given belt-pulley system with different data types. The detail operation of solving the differential equations is illustrate in subsection 5.2.1.

We hope to find a proper position of the tensioner providing a sufficient tension and achieving the maximum transmissive efficiency. A part of the belt-pulley system shows in Figure 5.1. The pulley<sub>1</sub> is the driver motor, pulley<sub>2</sub> is the tensioner, pulley<sub>3</sub> is the engine, and pulley<sub>4</sub> is the compressor of air conditioner. The optimization model and the engineering purposes of each constraint and objective function are demonstrated in subsection 5.2.2. The results of the optimization and the summary are shown in subsection 5.2.2 and 5.2.4 separately.



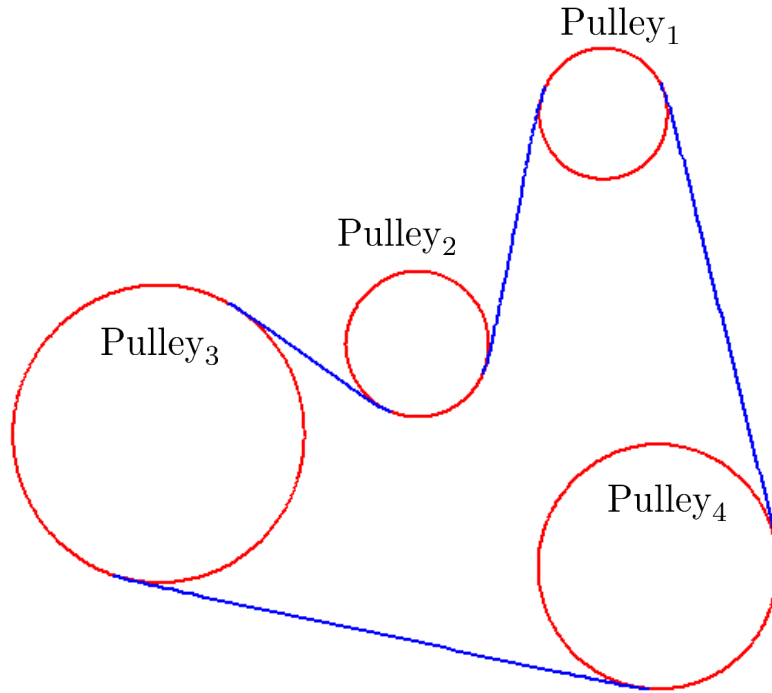


Figure 5.1: Two pulleys belt system

### 5.2.1 Numerical adjustment of belt-pulley system

The two basic governing in span zone is shown in Equation 5.5a.

$$(T - GV)' + EI\kappa\kappa' = 0$$

$$(T - GV)\kappa - EI\kappa'' = 0$$

From Equation 2.5, Let  $T - GV = W$  and we can yield three differential equations. For geometric requirements, other four differential equations are yielded. The  $s$  is the length of the span,  $0 \geq s \geq \hat{L}$

- Physical differential equations:

- $dW/ds = EI\kappa\kappa'$ , derived from Equation(5.5a).
- $d\kappa/ds = W/\kappa \times EI$ , derived from Equation(5.5a).
- $d^2\kappa/ds^2 = W\kappa/EI$ , derived from Equation(5.5a).

- Geometric differential equations:

- $d\theta/ds = \kappa$ , it is the definition of curvature.
- $dx/ds = \cos \theta$ ,  $x$  is a position on x-axis.
- $dy/ds = \sin \theta$ ,  $y$  is a position on y-axis.
- $d\hat{L}/ds = 0$ ,  $\hat{L}$  is the length of span in steady motion and it is a constant.

Some boundary conditions are needed to solve the physical and geometric differential equations. These boundaries are there to ensure the belt exactly contacts with pulleys and they are orthogonal at the points of contacts.

- Boundary conditions:

- $W(0) = T(0) - GV(0)$ , the initial value of span closed to pulley<sub>*i*</sub>
- $\kappa(0) = 1/R_i$ , at points of contacts, the curvature of belt is equal to which of pulleys.
- $\kappa(\hat{L}) = 1/R_{i+1}$
- $x(0) = R_i \sin(\theta(0))$ , ensuring the belt is orthogonal to pulleys.
- $x(\hat{L}) - D_i = R_{i+1} \sin(\theta(\hat{L}))$
- $x(0)^2 + y(0)^2 = R_i^2$ , ensuring the belt contact with pulleys.
- $x(\hat{L}) - D_i^2 + y(\hat{L})^2 = R_{i+1}^2$

There is a problem when solving these differential equations, the length of the span  $\hat{L}$  is unknown. We can solve this problem by setting these variables non-dimensional,  $\hat{s} = s/\hat{L}$ ,  $\hat{x} = x/\hat{L}$ ,  $\hat{y} = y/\hat{L}$ ,  $\hat{\kappa} = \kappa\hat{L}$ , and  $\hat{W} = W\hat{L}^2/EI$ . After this operation, the ranges of  $\hat{s}$  is normalized to  $0 \leq \hat{s} \leq 1$  Substitute these non-dimensional variables into seven differential and boundary conditions yields:

- Physical differential equations:

- $d\hat{W}/d\hat{s} = \hat{\kappa}\hat{\kappa}'$
- $d\hat{\kappa}/d\hat{s} = -\hat{W}/\hat{\kappa}$
- $d^2\hat{\kappa}/d\hat{s}^2 = \hat{W}\hat{\kappa}$

- Geometric differential equations:

- $d\theta/d\hat{s} = \kappa$
- $d\hat{x}/d\hat{s} = \cos \theta$
- $d\hat{y}/d\hat{s} = \sin \theta$
- $d\hat{L}/d\hat{s} = 0$

- Boundary conditions:

- $W(\hat{0}) = W(0)\hat{L}^2/EI$ , the initial value of span closed to pulley<sub>*i*</sub>
- $\kappa(\hat{0}) = \hat{L}/R_i$ , at points of contacts, the curvature of belt is equal to which of pulleys.
- $\kappa(1) = \hat{L}/R_{i+1}$
- $\hat{L}(0)\hat{x}(0) = R_i \sin(\theta(0))$ , ensuring the belt is orthogonal to pulleys.
- $\hat{L}\hat{x}(1) - D_i = R_{i+1} \sin(\theta(\hat{L}))$
- $\hat{L}\hat{x}(0)^2 + \hat{L}\hat{y}(0)^2 = R_i$ , ensuring the belt contact with pulleys.
- $\hat{L}\hat{x}(1) - D_i^2 + \hat{L}\hat{y}(1)^2 = R_{i+1}$

After non-dimensional operation, these seven differential equations are transferred into the standard boundary value problem(BVP). We can solve them by BVP solver, and obtain the solution about  $W$ ,  $\theta$ ,  $\kappa$ ,  $\kappa'$ ,  $\hat{L}$ ,  $x$  and  $y$ . In this thesis, we choose `bvp4c` as the solver in MATLAB toolbox. When using the `bvp4c`, the numerical initial values are required for each parameter. The improper initial values would lead a wrong results. The most unpredictable initial value is the length of span in the steady motion,  $\hat{L}$ . In order to proceed to the optimization, the initial value of  $\hat{L}$  has to be generate automatically. This initial value generator guess an initial value based on the distance between two pulleys. However, this generator does not work properly at each iteration, and the initial value of each iteration might change. To ensure the optimization work well, therefore, we display the maximum residual in the solving process and demand the residual is less the 0.1%, or re-solve the BVP with other initial values until the maximum residual satisfy the requirement.

## 5.2.2 The design model of belt-pulley systems

The properties of the belt such as bending stiffness,  $EI$ , the longitudinal stiffness,  $EA$ , and the fluctuation of radius of pulleys would influence the performance of the belt-pulley system

and they might deteriorate with time. Furthermore, other parameters do not vary with time but they would vary in a range due to the tolerance in the manufacture. We sample these uncertainties to design a reliability belt-pulley system that can work well over a period of time. There are three models to compare the optimal design with different quality of data

### Deterministic model

$$\begin{aligned}
 \min_x f(E(t), \mu(t), I) &= -\left(\sum_{i \neq 1} M_i \times \omega_i\right) M_1 \times \omega_1 \\
 \text{s.t. } g_1 &= (R_1 + R_2) - D_1 \leq 0 \\
 g_2 &= (R_2 + R_3) - D_2 \leq 0 \\
 g_3 &= (R_2 + R_4) - D_1 \leq 0 \\
 g_4(t) &= \max(T) - T_u \leq 0 \\
 g_5(t) &= T_l - \min(T) \leq 0
 \end{aligned} \tag{5.6}$$

where

$$\begin{cases}
 \mathbf{EI} = 0.05 \\
 \mathbf{EA} = 111200 \\
 R_2 = 40.75 \times 10^{-3} \\
 T_u = 10000 \\
 T_l = 0
 \end{cases}$$

### RBDO model

$$\begin{aligned}
\min_x f(E(t), \mu(t), I) &= -\left(\sum_{i \neq 1} M_i \times \omega_i\right) / M_1 \times \omega_1 \\
\text{s.t. } g_1 &= \Pr((R_1 + R_2) - D_1 \leq 0) \geq R_t \\
g_2 &= \Pr((R_2 + R_3) - D_2 \leq 0) \geq R_t \\
g_3 &= \Pr((R_2 + R_4) - D_1 \leq 0) \geq R_t \\
g_4(t) &= \Pr(\max(T) - T_u \leq 0) \geq R_t \\
g_5(t) &= \Pr(T_l - \min(T) \leq 0) \geq R_t \\
\text{where } &\begin{cases} \mathbf{EI} \sim N(0.05, 0.005) \\ \mathbf{EA} \sim N(111200, 100) \\ R_2 \sim N(40.75 \times 10^{-3}, 0.8) \\ T_u = 10000, T_l = 0 \\ R_t = 0.8 \end{cases}
\end{aligned} \tag{5.7}$$

### Bayesian RBDO model with life data

$$\begin{aligned}
\min_x f(E(t), \mu(t), I) &= -\left(\sum_{i \neq 1} M_i \times \omega_i\right) / M_1 \times \omega_1 \\
\text{s.t. } g_1 &= \Pr(\Pr((R_1 + R_2) - D_1 \leq 0) \geq R_t) \geq CR_t \\
g_2 &= \Pr(\Pr((R_2 + R_3) - D_2 \leq 0) \geq R_t) \geq CR_t \\
g_3 &= \Pr(\Pr((R_2 + R_4) - D_1 \leq 0) \geq R_t) \geq CR_t \\
g_4(t) &= \Pr(\Pr(\max(T) - T_u \leq 0) \geq R_t) \geq CR_t \\
g_5(t) &= \Pr(\Pr(T_l - \min(T) \leq 0) \geq R_t) \geq CR_t \\
g_6 &= \Pr(\Pr(T[n_{g_4}] > t) \geq P_t) \geq CL_t \\
g_7 &= \Pr(\Pr(T[n_{g_5}] > t) \geq P_t) \geq CL_t \\
\text{where } &\begin{cases} T_u = 10000, T_l = 0 \\ R_t = 0.8, CR_t = 0.8 \text{ and } CL_t = 0.8 \\ n_{g_4}, n_{g_5} = 3 \\ t = 10 \end{cases}
\end{aligned} \tag{5.8}$$

The first to third constraints in three models ensure the tensioner does not contact with pulleys and the fourth and fifth constraints are the maximum and minimum tension requirements. Furthermore, the sixth and seventh constraints in the Bayesian RBDO model govern

Table 5.4: Samples of the frictional bending stiffness  $EI(t)$

$EI(t = 0)$	0.0548	0.0527	0.0472	0.0513	0.0479
$EI(t = 1)$	0.0165	0.0086	0.0227	0.0137	0.0228
$EI(t = 2)$	0.0040	0.0040	-0.0015	0.0032	0.0074

Table 5.5: Samples of the longitudinal stiffness  $EA(t)$

$EA(t = 0)$	111328.29	111284.78	111150.55	111324.73	111174.16
$EA(t = 1)$	40772.73	40860.88	40911.12	40908.48	40968.29
$EA(t = 2)$	15101.02	15004.65	15000.58	15012.91	14996.11

the aging behavior of the fourth and fifth constraints. The failure number of the fourth and fifth constraints do not exceed 3 within 10 months.

### 5.2.3 Comparison of results and discussion

The design variables in the three optimization models are the  $x$  and  $y$  coordinate of the tensioner. The proper selection of tensioner position will have the maximum power efficiency of the belt-pulley system. In the each design step, the belt-pulley system would be analyzed with the a fixed tensioner position. In this section, we will discuss the performance of the belt-pulley system in the design process instead of just demonstrate the optimal solution of the three optimization models. We select two analyzing results which are yielded from with two position of the tensioner shown in the Figure 5.2. Table 5.7 is the objective function and the constraints of the each optimization model with the position of the tensioner,  $(x, y) = (0.1372, 0.0277)$ . Table 5.8 shows the objective function and the constraints of the each optimization model with the position of the tensioner,  $(x, y) = (0.2, 0.05)$ . By moving the position of the tensioner, we can observe the effects on the objective function and constraints.

In the deterministic optimization model with the first fixed position of the tensioner, only the 5th constraint is not satisfied and the minimum objective function is  $-0.0314$ , it means the power efficiency is about 3%. The analyzing result of RBDO model is much the

Table 5.6: Samples of the radius of pulley 2  $R_2$

$R_2(m)$	0.0300	0.0311	0.0303	0.0307	0.0306
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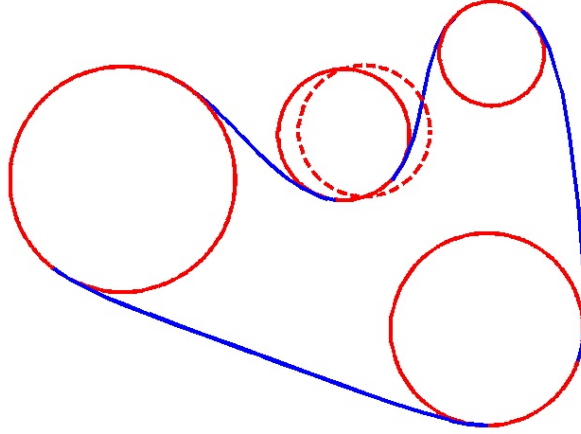


Figure 5.2: Two positions of the tensioner

same as in deterministic optimization model. The uncertainties of the  $EI$ ,  $EA$ , and  $R_2$  are modeled as normal distribution, and 1000 samples of each uncertainty are drawn from the normal distribution. Calculating the probability of the constraint  $g_1$  to  $g_5$  being satisfied and that probability has to be greater than the reliability target,  $R_t = 0.8$ . The 5th constraint in the RBDO model is not satisfied. In the Bayesian RBDO with life data, the deterioration of the material properties,  $EI(t)$  and  $EA(t)$ , are taken into consideration. The reliability targets and the confidence ranges and levels target are set as 0.8 and the requirement of aging constraints,  $g_6$  and  $g_7$ , are that the failure number of  $g_4$  and  $g_5$  do not exceed 3 within 10 months. The result of Bayesian RBDO with life data is different from the results of deterministic and RBDO model. The confidence ranges are not satisfied. However, the confidence ranges are not only influenced by the position of the tensioner but also by the size of the samples. The violated constraints in the Bayesian RBDO model do not exactly mean that the geometric requirements and the aging restrictions are not satisfied. Some constraints might be satisfied after adding new samples without moving the position of the tensioner.

With the second position of the tensioner,  $(x, y) = (0.2, 0.05)$ . The power efficiency in the deterministic model drops slightly. The first constraint is violated; that means the tensioner

Table 5.7: The analysis of the belt-pulley system with fixed position of tensioner  $(x, y) = (0.1372, 0.0277)$

Models	Objective function	Constraints
Deterministic	-0.03256	(-0.0314,-0.0292,-0.0493,-9312.5,1608.5)
RBDO	-0.03263	(-0.2,-0.2,-0.2,-0.16,0.8)
Bayesian RBDO with life data	-0.0629	(0.0776,0.0776,0.0776,0.999,0.876,0.999,-0.243)

Table 5.8: The analysis of the belt-pulley system with fixed position of tensioner  $(x, y) = (0.2, 0.05)$

Models	Objective function	Constraints
Deterministic	-0.03251	(0.0329,-0.0954,-0.0448,-10858,0)
RBDO	-0.03249	(0.8,-0.2,-0.2,-0.2,-0.2)
Bayesian RBDO with life data	0	(0.999,0.0776,0.0776,0.999,0.0777,0.999,-0.243)

interferes with pulleys. However, the maximum and minimum tension requirements are satisfied in this case. The minimum tension changes a lot with the position of the tensioner in the deterministic model. The result of RBDO model shows that the tensioner interferes with the pulley and the fifth constraint now is satisfied, and the value of objective function drop slightly too. The result of the Bayesian RBDO model has a similar tendency to deterministic and RBDO model, the constraints  $g_1$  and  $g_5$  is influenced most by the change of the position of the tensioner. However, the position of the tensioner does not influence the aging behavior greatly. In the Bayesian RBDO model, the aging process is estimated by the total failure number over a period of time and the number of subsection of time. The movement of the tensioner influence the total failure number only.

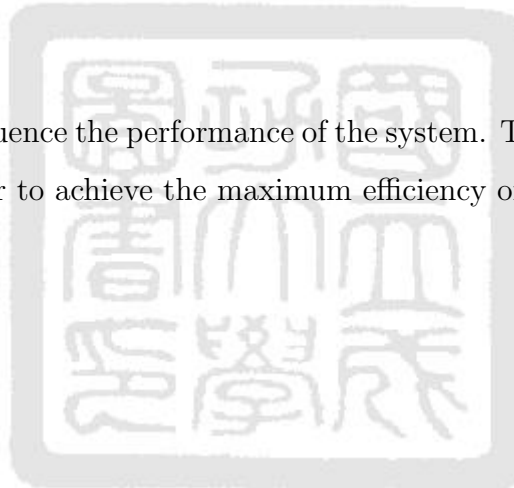


## 5.2.4 Summary

Overall, we have the following observations using the proposed method in the belt-pulley system.

1. With a fixed position of the tensioner, the result of the Bayesian RBDO might be influenced by the biased sample.
2. The aging constraints are not affected greatly by the change of the position of the tensioner.
3. In this belt-pulley system, the position of the tensioner influence the tension of the belt greatly.

The tension of the belt would influence the performance of the system. Therefore, we can choose a proper position of the tensioner to achieve the maximum efficiency of belt-pulley system.



# Chapter 6 Conclusion and Future Work

## 6.1 Conclusion

The main purpose of this thesis is to propose a design method considering the deterioration of components to address the time-dependent reliability-based design optimization with life data and satisfy the confidence and reliability targets.

In Chapter 3, the selection of the priors in the Bayesian inference is specified, and estimation of reliability of a constraints using Bayesian inference with life data is constructed and definition of the confidence ranges and confidence bounds is done. In Chapter 4, the overall design scheme is proposed. In this design scheme, the fewest additional sample is required. When adding new sample is necessary, the strategy of adding new samples is proposed based on the types of the constraints and the importance of the parameters. In Chapter 5, a mathematical example and a belt-pulley design show the validity of proposed design scheme.

Three main contributions of this thesis:

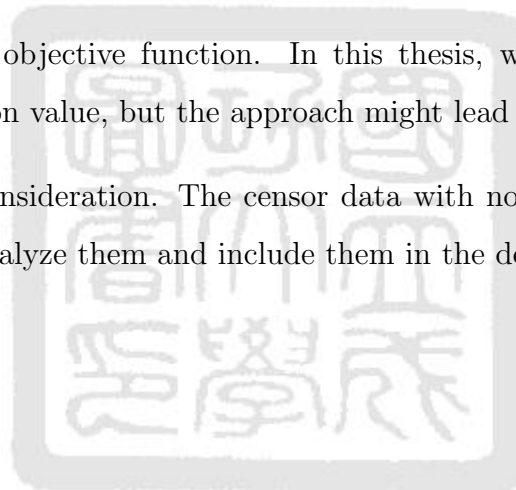
1. **Use the fewest number of sample to satisfy the confidence target and yield a optimal design under the confidence and reliability targets:** The additional sample is added to satisfy the lowest confidence target and in the design process, the optimization algorithm would force the design to satisfy the reliability target. If the reach the reliability target is impossible, the proposed method would tell the user there are no feasible solution with current size of sample.
2. **Allocate the resource to improve the confidence more efficiently:** The measurements of sample is expensive, therefore, the critical parameter is the only focus. For time-variant constraint, determine the critical parameter by analyzing the sensitivity of the time-dependent parameter; for time-invariant constraints, analyze the sensitivity of all parameters, and define the most sensitive parameter as the critical parameter.
3. **Make the BVP solver more steady while optimization:** The optimization process would be terminated by encountering the singular point while solving the BVP. We

construct a initial value generator to choose a proper initial value to avoid the singular points. While encounter the singular points, we force the optimization process continue, but the ill solution will be excluded.

## 6.2 Future work

There are some points deserving further investigation:

1. The strategy of modifying the reliability targets, while the optimization target can not be satisfy under current size of samples.
2. The impact of sample on objective function. In this thesis, we simply minimize the maximum objective function value, but the approach might lead too strict result.
3. Taking censor data into consideration. The censor data with no exact failure time and value need other way to analyze them and include them in the design method.



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# Personal Communication

姓名: 韓侑君

歷年學歷:

新竹女子高級中學(2003/6-2006/9)

成功大學機械工程學系(2006/9-2011/6)

成功大學機械所 (2011/9-2013/6)

聯絡信箱: [falldancing@hotmail.com](mailto:falldancing@hotmail.com)

